

Mitigating Multicollinearity in Linear Regression Model with Two Parameter Kibria-Lukman Estimators

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Abstract: - This study delves into the challenges faced by the ordinary least square (OLS) estimator, traditionally regarded as the Best Linear Unbiased Estimator in classical linear regression models. Despite its reliability under specific conditions, OLS falters in the face of multicollinearity, a problem frequently encountered in regression analyses. To combat this issue, various ridge regression estimators have been developed, characterized as one-parameter and two-parameter ridge-type estimators. In this context, our research introduces novel two-parameter estimators, building upon a recently developed one-parameter ridge estimator to mitigate the impact of multicollinearity in linear regression models. Theoretical analysis and simulation experiments were conducted to assess the performance of the proposed estimators. Remarkably, our results reveal that, under certain conditions, these new estimators outperform existing estimators, displaying a significantly reduced mean square error. To validate these findings, real-life data was employed, aligning with the outcomes derived from theoretical analysis and simulations.

Key-Words: - Kibria-Lukman Estimator, New Two-parameter Kibria-Lukman Estimator, Modified Two-parameter Kibria-Lukman Estimator, Monte Carlo Simulation, Multicollinearity, Mean Square Error.

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1 Introduction

Consider the general linear regression model defined in matrix form as:

$$y = X\beta + \varepsilon, \quad (1)$$

where y is a $n \times 1$ vector of the response variable, X is a known $n \times p$ full-rank matrix of predictor variables, β is a $p \times 1$ vector of unknown regression parameters to be estimated, and ε is $n \times 1$ vector of random error such that $E(\varepsilon) = 0$ and $Cov(\varepsilon) = \sigma^2 I$. Equation (1) can be written in a canonical form as:

$$y = Z\alpha + \varepsilon \quad (2)$$

where $Z = XQ$, $\alpha = Q'\beta$ and Q is the orthogonal matrix whose columns constitute the eigenvectors of XX' . Then $Z'Z = Q'XXQ = \Lambda = \text{diag}(\lambda_1, \dots, \lambda_p)$,

where $\lambda_1 \geq \lambda_2 \geq \dots \lambda_p > 0$ are the ordered eigen values XX' . The ordinary least square estimator (OLSE) of β in (1) can be defined as:

$$\hat{\alpha}_{OLS} = \Lambda^{-1} X'y \quad (3)$$

where $\Lambda = XX'$ is the design matrix.

The Ordinary Least Squares (OLS) estimator is considered the Best Linear Unbiased Estimator (BLUE) when all assumptions of the classical linear regression model remain intact and unviolated. This characteristic has established the OLS estimator as the most powerful and popular tool for estimating regression models, [1]. To utilize the OLS estimator for estimating model parameters in linear regression, one essential assumption is the non-correlation of explanatory variables. However, the problem of multicollinearity arises when explanatory variables exhibit strong correlations or linear relationships, [2]. Multicollinearity, if present, severely affects the OLS estimator, leading to unstable and imprecise parameter estimates, questionable predictions, and invalid statistical inferences about model parameters. It also results in regression coefficients with exaggerated absolute values and signs that can reverse with minor changes in the data, [3], [4]. It also impacts t-tests, the extra sum of squares, fitted values, predictions, and coefficient of determination, [5].

To address the issue of multicollinearity, various biased estimators have been developed. Some of

these one-parameter estimators include the Stein estimator [6], principal component estimator [7], ordinary ridge regression estimator by [8], modified ridge regression by [9], contraction estimator [10], Liu estimator [11], and KL [12]. Additionally, certain authors have introduced two-parameter estimators to combat multicollinearity, such as [13], [14], [15], [16], [17], [18], [19] and [20]. The ordinary ridge regression estimator (ORRE) is one of the most widely used biased estimators. It intends to overcome the multicollinearity problem by adding a positive value, biasing or shrinkage parameter k , to the ill-conditioned matrix's diagonal elements XX' . The selection of k is a major problem of ORRE. This is because the biasing parameter k plays a significant role in controlling the regression's bias toward the mean of the dependent variable, [8].

The Ordinary ridge regression estimator, as proposed by [8], stands as one of the most widely used among these estimators. It effectively addresses the issue of multicollinearity by introducing a positive value, denoted as k , to the diagonal elements of the $Z'Z$ matrix. This constant k serves as the biasing parameter. However, a significant challenge associated with ridge regression lies in selecting this biasing parameter k , as it plays a crucial role in controlling the regression's bias towards the mean of the dependent variable. Various authors have put forth proposals for the biasing parameter k . Some notable contributors include, [21], [22], [23], [24], [25], [26], [27], [28], [29], [30] and numerous others. [11], proposed another estimator with d as the biasing parameter. The Liu Estimator gains preference due to its capacity for appropriate d selection, being a linear function of the biasing parameter d . However, selecting a suitable k remains challenging because the ordinary ridge regression estimator relies on a nonlinear function of the biasing parameter k [8].

This paper aims to introduce new two-parameter estimators for regression parameter estimation in scenarios where independent variables exhibit correlation. The performance of the proposed estimators is compared with the OLSE, ridge regression [8], Liu estimator [11], two-parameter estimator (TP) [13], Modified Ridge Type (MRT) [17], and Kibria-Lukman (KL) [12].

2 Some Alternative Biased Estimators and the Proposed Estimator

2.1 Some Estimators as Alternative to OLS

Ridge-type estimators have been proposed as an alternative to the OLSE. The canonical form of OLSE is written in equation (3). Following this, the ordinary ridge regression (RE) proposed by [8] is given as:

$$\hat{\alpha}_{RE}(k) = (\Lambda + kI)^{-1} Z' y = H_0^{-1} Z' y \quad (4)$$

where k is the non-negative constant known as the biasing parameter.

The Liu estimator (LE) [11] is defined as:

$$\hat{\alpha}_{LE}(d) = [\Lambda + I]^{-1} [\Lambda + dI] \hat{\alpha}_{OLS} = D_0 D_1 \hat{\alpha}_{OLS} \quad (5)$$

where d is the biasing parameter of the Liu Estimator.

The KL estimator [12] is given as:

$$\hat{\alpha}_{KL}(k) = (\Lambda + kI)^{-1} (\Lambda - kI) \hat{\alpha}_{OLS} = H_0^{-1} H_1 \hat{\alpha}_{OLS} \quad (6)$$

TP estimator [13] is given as:

$$\hat{\alpha}_{TP}(k, d) = (\Lambda + kI)^{-1} (\Lambda + kdI) \hat{\alpha}_{OLS} = H_0^{-1} H_2 \hat{\alpha}_{OLS} \quad (7)$$

where k and d are the biasing parameters of the ridge and Liu Estimator, respectively.

The Modified Ridge Type (MRT) Parameters proposed [17] is given as:

$$\hat{\alpha}_{MRT}(k, d) = (\Lambda + k(1+d)I)^{-1} \Lambda \hat{\alpha}_{OLS} = R_k \hat{\alpha}_{OLS} \quad (8)$$

where $H_0 = (\Lambda + kI)$, $H_1 = (\Lambda - kI)$, $H_2 = (\Lambda + kdI)$, $H_3 = (\Lambda - kdI)$, $D_0 = (\Lambda + I)^{-1}$, $D_1 = (\Lambda + dI)$ and $R_k = \Lambda(\Lambda + k(1+d)I)^{-1}$.

2.2 New Two parameter Kibria-Lukman (NTPKL) Estimator

In this article, the ridge estimator was grafted into the Kibria-Lukman estimator, [12], to propose new parameter estimators. The first proposed two-parameter estimator is defined as follows:

$$\hat{\alpha}_{NTPKL}(k, d) = (\Lambda + kd_1 I)^{-1} (\Lambda - kd_2 I) \hat{\alpha}_{RE} \quad (9)$$

$$\hat{\alpha}_{NTPKL}(k, d) = H_2^{-1} H_3 \hat{\alpha}_{RE} \quad (10)$$

where $k > 0$, $0 < d < 1$ and $d_1 = d_2$

The following are the properties of the NTPKL Estimator

$$E(\hat{\alpha}_{NTPKL}(k, d)) = \Lambda H_2^{-1} H_3 H_0^{-1} \alpha \quad (11)$$

$$B(\hat{\alpha}_{NTPKL}(k, d)) = (\Lambda H_2^{-1} H_3 H_0^{-1} - I) \alpha \quad (12)$$

$$D[\hat{\alpha}_{NTPKL}(k, d)] = \sigma^2 H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1} \quad (13)$$

The Mean Square Error Matrix (MSEM) of the proposed estimator is given as:

$$MSEM[\hat{\alpha}_{NTPKL}(k, d)] = \sigma^2 H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1} + (\Lambda H_2^{-1} H_3 H_0^{-1} - I) \alpha \alpha' (\Lambda H_2^{-1} H_3 H_0^{-1} - I)' \quad (14)$$

2.3 Modified Two Parameter Kibria-Lukman (MTPKL) Estimators

The Modified Two Parameter Kibria-Lukman Estimators are special cases of the New Two Parameter Kibra-Lukman Estimator in equation (9). The first Modified estimator (MTPKL1) is obtained by equating d_2 to one (i.e., $d_2 = 1$) in equation (9).

$$\hat{\alpha}_{MTPKL1}(k, d) = (\Lambda + kdI)^{-1} (\Lambda - kI) \hat{\alpha}_{RE} \quad (15)$$

$$\hat{\alpha}_{MTPKL1}(k, d) = H_2^{-1} H_1 \hat{\alpha}_{RE} \quad (16)$$

where $k > 0$ and $0 < d < 1$.

The following are the properties of the MTPKL1 Estimator:

$$E(\hat{\alpha}_{MTPKL1}(k, d)) = \Lambda H_0^{-1} H_1 H_2^{-1} \alpha \quad (17)$$

$$B(\hat{\alpha}_{MTPKL1}(k, d)) = (\Lambda H_0^{-1} H_1 H_2^{-1} - I) \alpha \quad (18)$$

$$D[\hat{\alpha}_{MTPKL1}(k, d)] = \sigma^2 H_0^{-1} H_1 H_2^{-1} \Lambda H_0^{-1} H_1 H_2^{-1} \quad (19)$$

The Mean Square Error Matrix (MSEM) of the proposed MTPKL1 estimator is given as:

$$MSEM[\hat{\alpha}_{MTPKL1}(k, d)] = \sigma^2 H_0^{-1} H_1 H_2^{-1} \Lambda H_0^{-1} H_1 H_2^{-1} + (\Lambda H_0^{-1} H_1 H_2^{-1} - I) \alpha \alpha' (\Lambda H_0^{-1} H_1 H_2^{-1} - I)' \quad (20)$$

The second Modified estimator (MTPKL2) is obtained by equating d_1 to one (i.e., $d_1 = 1$) in equation (9). It was also obtained by modifying, [13], following the idea of the $(k - d)$ class estimator proposed by [31]. Thus, the MTPKL2 estimator can be defined as:

$$\hat{\alpha}_{MTPKL2}(k, d) = (\Lambda + kI)^{-1} (\Lambda - kdI) \hat{\alpha}_{RE} \quad (21)$$

$$\hat{\alpha}_{MTPKL2}(k, d) = H_0^{-1} H_3 \hat{\alpha}_{RE} \quad (22)$$

where $k > 0$ and $0 < d < 1$.

The following are the properties of the MTPKL2 Estimator:

$$E(\hat{\alpha}_{MTPKL2}(k, d)) = \Lambda H_0^{-1} H_3 H_0^{-1} \alpha \quad (23)$$

$$B(\hat{\alpha}_{MTPKL2}(k, d)) = (\Lambda H_0^{-1} H_3 H_0^{-1} - I) \alpha \quad (24)$$

$$D[\hat{\alpha}_{MTPKL2}(k, d)] = \sigma^2 H_0^{-1} H_3 H_0^{-1} \Lambda H_0^{-1} H_3 H_0^{-1} \quad (25)$$

The Mean Square Error Matrix (MSEM) of the proposed MTPKL2 estimator is given as:

$$\begin{aligned} MSEM[\hat{\alpha}_{MTPKL2}(k, d)] &= \sigma^2 H_0^{-1} H_3 H_0^{-1} \Lambda H_0^{-1} H_3 H_0^{-1} \\ &+ (\Lambda H_0^{-1} H_3 H_0^{-1} - I) \alpha \alpha' (\Lambda H_0^{-1} H_3 H_0^{-1} - I)' \end{aligned} \quad (26)$$

Since the MTPKL1 and MTPKL2 estimators are special cases of NTPKL, Lemmas 1, 2 and 3 will be used to make some theoretical comparisons between NTPKL and six existing estimators and to prove the statistical properties of the NTPKL Estimator.

Lemma 1. Let $n \times n$ matrices $M > 0$ and $N > 0$ (or $N \geq 0$), Then $M > N$ if and only if $\lambda_i(NM^{-1}) < 1$ where $\lambda_i(NM^{-1})$ is the largest eigenvalue of the matrix NM^{-1} , [32].

Lemma 2. Let M be an $n \times n$ positive definite matrix, that is, $M > 0$, and α be some vector, then, $M - \alpha \alpha' \geq 0$ if and only if $\alpha' M^{-1} \alpha \leq 1$, [33].

Lemma 3. Let $\hat{\alpha}_i = A_i y$, $i = 1, 2$, be two linear estimators of α . Suppose that $D = Cov(\hat{\alpha}_1) - Cov(\hat{\alpha}_2) > 0$, where $Cov(\hat{\alpha}_i)$, $i = 1, 2$ denotes the covariance matrix of $\hat{\alpha}_i$ and $b_i = Bias(\hat{\alpha}_i) = (A_i X - I) \alpha$, $i = 1, 2$. Consequently,

$$\begin{aligned} \Delta(\hat{\alpha}_1 - \hat{\alpha}_2) &= MSEM(\hat{\alpha}_1) - MSEM(\hat{\alpha}_2) \\ &= \sigma^2 D + b_1 b_1' - b_2 b_2' > 0 \end{aligned} \quad (27)$$

if and only if

$$b_2' [\sigma^2 D + b_1 b_1']^{-1} b_2 < 1 \quad [34]$$

where $MSEM(\hat{\alpha}_i) = Cov(\hat{\alpha}_i) + b_i b_i'$

2.4 Comparison of NTPKL Estimator with Existing Estimators

In this section, the theoretical comparison is carried out among the estimators to examine the performance of the proposed estimator, $\hat{\alpha}_{NTPKL}(k, d)$ over other estimators; $\hat{\alpha}_{OLS}$, $\hat{\alpha}_{RE}$, $\hat{\alpha}_{LE}$, $\hat{\alpha}_{NTP}$, $\hat{\alpha}_{MRT}$, $\hat{\alpha}_{KL}$.

2.4.1 Comparison between $\hat{\alpha}_{OLS}$ and $\hat{\alpha}_{NTPKL}(k, d)$

The MSEM of the estimator $\hat{\alpha}_{OLS} = \Lambda^{-1} Z' y$ is as follows:

$$MSEM[\hat{\alpha}_{OLS}] = \sigma^2 \Lambda^{-1} \quad (28)$$

$$\begin{aligned} \text{The difference between (14) and (28)} \\ MSEM(\hat{\alpha}_{OLS}) - MSEM[\hat{\alpha}_{NTPKL}(k, d)] &= \sigma^2 \Lambda^{-1} - \sigma^2 H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1} \\ &+ (\Lambda H_2^{-1} H_3 H_0^{-1} - I) \alpha \alpha' (\Lambda H_2^{-1} H_3 H_0^{-1} - I)' \end{aligned} \quad (29)$$

Let $k > 0$ and $0 < d < 1$. Thus, the following theorem holds.

Theorem: The proposed estimator $\hat{\alpha}_{NTPKL}(k, d)$ is superior to $\hat{\alpha}_{OLS}$ if and only if $\alpha' (\Lambda H_2^{-1} H_3 H_0^{-1} - I)' \sigma^2 [(\Lambda^{-1} - H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1})]^{-1} (\Lambda H_2^{-1} H_3 H_0^{-1} - I) \alpha < 1$ (30)

Proof

$$\begin{aligned} D(\hat{\alpha}_{OLS}) - D(\hat{\alpha}_{NTPKL}(k, d)) &= \\ \sigma^2 (\Lambda^{-1} - H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1}) & \\ = \sigma^2 \text{diag} \left\{ \frac{1}{\lambda_i} - \frac{\lambda_i (\lambda_i - kd)^2}{(\lambda_i + kd)^2 (\lambda_i + k)^2} \right\}_{i=1}^p \end{aligned} \quad (31)$$

$\Lambda^{-1} - H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1}$ will be pdf if and only if $(\lambda_i + kd)^2 (\lambda_i + k)^2 - \lambda_i^2 (\lambda_i - kd)^2 > 0$. By lemma 3, the proof is completed.

2.4.2 Comparison between $\hat{\alpha}_{RE}(k)$ and $\hat{\alpha}_{NTPKL}(k, d)$

The bias vector, covariance matrix, and MSEM of the estimator $\hat{\alpha}_{RE}(k) = (\Lambda + kI)^{-1} Z' y$ are as follows:

$$E[\hat{\alpha}_{RE}(k)] = (\Lambda + kI)^{-1} \Lambda \alpha \quad (32)$$

$$D[\hat{\alpha}_{RE}(k)] = \sigma^2 (\Lambda + kI)^{-1} \Lambda (\Lambda + kI)^{-1} \quad (33)$$

and the Mean Square Error Matrix is given as: $MSEM[\hat{\alpha}_{RE}(k)] = \sigma^2 H_0^{-1} \Lambda H_0^{-1} + k^2 H_0^{-1} \alpha \alpha' H_0^{-1}$ (34)

The difference between (14) and (34)

$$\begin{aligned} &MSEM(\hat{\alpha}_{RE}(k)) - MSEM[\hat{\alpha}_{NTPKL}(k, d)] \\ &= \sigma^2 H_0^{-1} \Lambda H_0^{-1} - \sigma^2 H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1} + k^2 H_0^{-1} \alpha \alpha' H_0^{-1} \\ &\quad - (\Lambda H_2^{-1} H_3 H_0^{-1} - I) \alpha \alpha' (\Lambda H_2^{-1} H_3 H_0^{-1} - I)' \end{aligned} \quad (35)$$

Let $k > 0$ and $0 < d < 1$. Thus, the following theorem holds.

Theorem: The proposed estimate $\hat{\alpha}_{NTPKL}(k, d)$ is superior to $\hat{\alpha}_{RE}(k)$ if and only if

$$\begin{aligned} &MSEM[\hat{\alpha}_{RE}(k)] - MSEM[\hat{\alpha}_{NTPKL}(k, d)] > 0 \text{ if and} \\ &\text{only if} \\ &\alpha' (\Lambda H_2^{-1} H_3 H_0^{-1} - I)' \end{aligned}$$

$$\begin{aligned} &\left[\sigma^2 (H_0^{-1} \Lambda H_0^{-1} - H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1}) + k^2 H_0^{-1} \alpha \alpha' H_0^{-1} \right]^{-1} \\ &(\Lambda H_2^{-1} H_3 H_0^{-1} - I) \alpha < 1 \end{aligned} \quad (36)$$

Proof: Considering the dispersion matrix difference between $D[\hat{\alpha}_{RE}(k)]$ and $D[\hat{\alpha}_{NTPKL}(k, d)]$

$$\begin{aligned} D_d &= \sigma^2 H_0^{-1} \Lambda H_0^{-1} - \sigma^2 H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1} \\ D_d &= \sigma^2 (\Lambda + kI)^{-1} \Lambda (\Lambda + kI)^{-1} - \sigma^2 (\Lambda + kdI)^{-1} (\Lambda - kdI) \\ &(\Lambda + kI)^{-1} \Lambda (\Lambda + kdI)^{-1} (\Lambda - kdI) (\Lambda + kI)^{-1} \\ D_d &= \sigma^2 (\Lambda + kI)^{-1} (\Lambda + kdI)^{-1} \\ &\left[\Lambda (\Lambda + kdI)^2 - \Lambda (\Lambda - kdI)^2 \right] (\Lambda + kI)^{-1} (\Lambda + kdI)^{-1} \\ D_d &= \sigma^2 (\Lambda + kI)^{-1} (\Lambda + kdI)^{-1} \\ &\left[4\Lambda^2 kdI \right] (\Lambda + kI)^{-1} (\Lambda + kdI)^{-1} \end{aligned} \quad (37)$$

It is observed that D_d is positive definite. By lemma 3, the proof is completed.

2.4.3 Comparison between $\hat{\alpha}_{LE}(d)$ and $\hat{\alpha}_{NTPKL}(k, d)$

The bias vector, covariance matrix, and MSEM of the estimator $\hat{\alpha}_{Liu} = (\Lambda + I)^{-1} (\Lambda + dI) \hat{\alpha}_{OLS}$ are as follows:

$$B[\hat{\alpha}_{LE}(d)] = (d-1)(\Lambda + I)^{-1} \alpha \quad (38)$$

$$D[\hat{\alpha}_{LE}] = \sigma^2 (\Lambda + I)^{-1} (\Lambda + dI)^{-1} \Lambda^{-1} (\Lambda + I)^{-1} (\Lambda + dI)^{-1} \quad (39)$$

$$\begin{aligned} MSEM[\hat{\alpha}_{LE}] &= \sigma^2 D_0 D_1 \Lambda^{-1} D_0' D_1' + \\ &(d-1) D_0 \alpha \alpha' (d-1)' D_0' \end{aligned} \quad (40)$$

Theorem: The proposed estimator $\hat{\alpha}_{NTPKL}(k, d)$ is superior to $\hat{\alpha}_{LE}(d)$ if and only if

$MSEM[\hat{\alpha}_{LE}(d)] - MSEM[\hat{\alpha}_{NTPKL}(k, d)] > 0$ if and only if

$$\begin{aligned} &\alpha' (\Lambda H_2^{-1} H_3 H_0^{-1} - I)' \left[\sigma^2 \left(\begin{aligned} &D_0 D_1 \Lambda^{-1} D_0' D_1' - \\ &\sigma^2 H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1} \end{aligned} \right) + \right. \\ &\left. (d-1) D_0 \alpha \alpha' (d-1)' D_0' \right]^{-1} \\ &(\Lambda H_2^{-1} H_3 H_0^{-1} - I) \alpha < 1 \end{aligned}$$

Proof: Considering the dispersion matrix difference between $D[\hat{\alpha}_{LE}(d)]$ and $D[\hat{\alpha}_{NTPKL}(k, d)]$

$$\begin{aligned} D_d &= \sigma^2 D_0 D_1 \Lambda^{-1} D_0' D_1' - \sigma^2 H_2^{-1} H_3 H_0^{-1} \Lambda H_2^{-1} H_3 H_0^{-1} \\ D_d &= \sigma^2 (\Lambda + I)^{-1} (\Lambda + dI)^{-1} \Lambda^{-1} (\Lambda + I)^{-1} (\Lambda + dI)^{-1} - \\ &\sigma^2 (\Lambda + kdI)^{-1} (\Lambda - kdI) (\Lambda + kdI)^{-1} \Lambda \\ &(\Lambda + kdI)^{-1} (\Lambda - kdI) (\Lambda + kdI)^{-1} \\ D_d &= \sigma^2 \text{diag} \left\{ \frac{(\lambda_i + d)^2}{\lambda_i (\lambda_i + 1)^2} - \frac{\lambda_i (\lambda_i - kd)^2}{(\lambda_i + kd)^2 (\lambda_i + k)^2} \right\}_i \end{aligned} \quad (41)$$

will be pdf if and only if $(\lambda_i + d)^2 (\lambda_i + kd)^2 (\lambda_i + k)^2 - \lambda_i^2 (\lambda_i - kd)^2 (\lambda_i + 1)^2 > 0$. For $0 < d < 1$ and $k > 0$, it was observed that $(\lambda_i + d)^2 (\lambda_i + kd)^2 (\lambda_i + k)^2 - \lambda_i^2 (\lambda_i - kd)^2 (\lambda_i + 1)^2 > 0$

By lemma 3, the proof is completed.

2.4.4 Comparison between $\hat{\alpha}_{KL}(k)$ and $\hat{\alpha}_{NTPKL}(k, d)$

The bias vector, covariance matrix, and MSEM of the estimator $\hat{\alpha}_{KL}(k) = (\Lambda + kI)^{-1}(\Lambda - kI)\hat{\alpha}$ are as follows:

$$B[\hat{\alpha}_{KL}(k)] = [H_0^{-1}H_1 - I]\alpha \quad (42)$$

$$D[\hat{\alpha}_{KL}(k)] = \sigma^2 H_0^{-1}H_1\Lambda^{-1}H_0^{-1}H_1' \quad (43)$$

$$MSEM[\hat{\alpha}_{KL}(k)] = \sigma^2 H_0^{-1}H_1\Lambda^{-1}H_0^{-1}H_1 + [H_0^{-1}H_1 - I]\alpha\alpha'[H_0^{-1}H_1 - I]' \quad (44)$$

Theorem: The proposed estimator $\hat{\alpha}_{NMKL2}(k, d)$ is superior to $\hat{\alpha}_{KL}(k)$ if and only if:

$MSEM[\hat{\alpha}_{KL}(k)] - MSEM[\hat{\alpha}_{NTPKL}(k, d)] > 0$ if and only if

$$\alpha'(\Lambda H_2^{-1}H_3H_0^{-1} - I) \left[\sigma^2 \begin{pmatrix} H_0^{-1}H_1\Lambda^{-1}H_0^{-1}H_1 - \\ H_2^{-1}H_3H_0^{-1}\Lambda H_2^{-1}H_3H_0^{-1} \\ + H_0^{-1}H_1\alpha\alpha'H_0^{-1}H_1 \end{pmatrix} \right]^{-1} (\Lambda H_2^{-1}H_3H_0^{-1} - I)\alpha < 1 \quad (45)$$

Proof: Considering the dispersion matrix difference between $D[\hat{\alpha}_{KL}(k)]$ and $D[\hat{\alpha}_{NTPKL}(k, d)]$

$$D_d = \sigma^2 H_0^{-1}H_1\Lambda^{-1}H_0^{-1}H_1 - \sigma^2 H_2^{-1}H_3H_0^{-1}\Lambda H_2^{-1}H_3H_0^{-1} \\ D_d = \sigma^2 (\Lambda - kI)(\Lambda + kI)^{-1} \Lambda^{-1}(\Lambda - kI)(\Lambda + kI)^{-1} - \sigma^2 (\Lambda + kdI)^{-1}(\Lambda - kdI)(\Lambda + kI)^{-1} \Lambda(\Lambda + kdI)^{-1} \\ (\Lambda - kdI)(\Lambda + kI)^{-1}$$

$$D_d = \sigma^2 \text{diag} \left\{ \frac{(\lambda_i - k)^2}{\lambda_i(\lambda_i + k)^2} - \frac{\lambda_i(\lambda_i - kd)^2}{(\lambda_i + kd)^2(\lambda_i + k)^2} \right\}_i^p$$

will be pdf if and only if $(\lambda_i - k)^2(\lambda_i + kd)^2 - \lambda_i^2(\lambda_i - kd)^2 > 0$. For $0 < d < 1$ and $k > 0$, it was observed that $(\lambda_i - k)^2(\lambda_i + kd)^2 - \lambda_i^2(\lambda_i - kd)^2 > 0$. By lemma 3, the proof is completed.

2.4.5 Comparison between $\hat{\alpha}_{MRT}(k, d)$ and $\hat{\alpha}_{NTPKL}(k, d)$

The bias vector, covariance matrix, and MSEM of the estimator $\hat{\alpha}_{MRT}(k, d) = (\Lambda + k(1+d)I)^{-1}\Lambda\hat{\alpha}_{OLS} = R_k\hat{\alpha}_{OLS}$ are as follows:

$$B[\hat{\alpha}_{MRT}(k, d)] = [R_k - I]\alpha \quad (46)$$

$$D[\hat{\alpha}_{MRT}(k, d)] = \sigma^2 R_k \Lambda^{-1} R_k' \quad (47)$$

$$MSEM[\hat{\alpha}_{MRT}(k, d)] = \sigma^2 R_k \Lambda^{-1} R_k' + [R_k - I]\alpha\alpha'[R_k - I]' \quad (48)$$

Where $R_k = \Lambda(\Lambda + k(1+d)I)^{-1}$. Let $k > 0$ and $0 < d < 1$. Thus, the following theorem holds.

Theorem: The proposed estimator $\hat{\alpha}_{NTPKL}(k, d)$ is superior to $\hat{\alpha}_{MRT}(k, d)$ if and only if

$MSEM[\hat{\alpha}_{MRT}(k, d)] - MSEM[\hat{\alpha}_{NTPKL}(k, d)] > 0$ if and only if

$$\alpha'(\Lambda H_2^{-1}H_3H_0^{-1} - I) \left[\sigma^2 (R_k \Lambda^{-1} R_k' - \sigma^2 H_2^{-1}H_3H_0^{-1}\Lambda H_2^{-1}H_3H_0^{-1}) + R_k\alpha\alpha'R_k' \right]^{-1} (\Lambda H_2^{-1}H_3H_0^{-1} - I)\alpha < 1 \quad (49)$$

Proof: Considering the dispersion matrix difference between $D[\hat{\alpha}_{MRT}(k, d)]$ and $D[\hat{\alpha}_{NTPKL}(k, d)]$

$$D_d = \sigma^2 R_k \Lambda^{-1} R_k' - \sigma^2 H_2^{-1}H_3H_0^{-1}\Lambda H_2^{-1}H_3H_0^{-1} \\ D_d = \sigma^2 (\Lambda + k(1+d)I)^{-1} \Lambda(\Lambda + k(1+d)I)^{-1} - \sigma^2 (\Lambda + kdI)^{-1}(\Lambda - kdI)(\Lambda + kI)^{-1} \Lambda(\Lambda + kdI)^{-1} \\ (\Lambda - kdI)(\Lambda + kI)^{-1}$$

$$D_d = \sigma^2 \text{diag} \left\{ \frac{\lambda_i}{(\lambda_i + k(1+d))^2} - \frac{\lambda_i(\lambda_i - kd)^2}{(\lambda_i + kd)^2(\lambda_i + k)^2} \right\}_{i=1}^p \quad (50)$$

will be pdf if and only if $\lambda_i(\lambda_i + kd)^2(\lambda_i + k)^2 - \lambda_i(\lambda_i - kd)^2(\lambda_i + k(1+d))^2 > 0$. For $0 < d < 1$ and $k > 0$, it was observed that

$$\lambda_i(\lambda_i + kd)^2(\lambda_i + k)^2 - \lambda_i(\lambda_i - kd)^2(\lambda_i + k(1+d))^2 > 0.$$

By lemma 3, the proof is completed.

2.4.6 Comparison between $\hat{\alpha}_{TP}(k, d)$ and $\hat{\alpha}_{NTPKL}(k, d)$

The bias vector, covariance matrix, and MSEM of the estimator

$\hat{\alpha}_{TP}(k, d) = (\Lambda + kI)^{-1}(\Lambda + kdI)\hat{\alpha}_{OLS}$ are as follows:

$$B[\hat{\alpha}_{TP}(k, d)] = [H_0^{-1}H_2 - I]\alpha \tag{51}$$

$$D[\hat{\alpha}_{TP}(k, d)] = \sigma^2 H_0^{-1}H_2\Lambda^{-1}H_0^{-1}H_2' \tag{52}$$

$$MSEM[\hat{\alpha}_{TP}(k, d)] = \sigma^2 H_0^{-1}H_2\Lambda^{-1}H_0^{-1}H_2 + [H_0^{-1}H_2 - I]\alpha\alpha'[H_0^{-1}H_2 - I]' \tag{53}$$

Let $k > 0$ and $0 < d < 1$. Thus, the following theorem holds.

Theorem: The proposed estimator $\hat{\alpha}_{NTPKL}(k, d)$ is superior to $\hat{\alpha}_{TP}(k, d)$ if and only if

$$MSEM[\hat{\alpha}_{TP}(k, d)] - MSEM[\hat{\alpha}_{NTPKL}(k, d)] > 0.$$

That is, if and only if,

$$\alpha'(\Lambda H_2^{-1}H_3H_0^{-1} - I)' \left[\sigma^2 \begin{pmatrix} H_0^{-1}H_2\Lambda^{-1}H_0^{-1}H_2 \\ -H_2^{-1}H_3H_0^{-1}\Lambda H_2^{-1}H_0^{-1} \end{pmatrix} + \begin{pmatrix} (H_0^{-1}H_2 - I)\alpha\alpha'(H_0^{-1}H_2 - I)' \\ (H_0^{-1}H_3 - I)\alpha\alpha'(H_0^{-1}H_3 - I)' \end{pmatrix} \right]^{-1} (\Lambda H_2^{-1}H_3H_0^{-1} - I)\alpha < 1 \tag{54}$$

Proof: Considering the dispersion matrix difference between $D[\hat{\alpha}_{TP}(k, d)]$ and $D[\hat{\alpha}_{NTPKL}(k, d)]$

$$D_d = \sigma^2 H_0^{-1}H_2\Lambda^{-1}H_0^{-1}H_2 - \sigma^2 H_2^{-1}H_3H_0^{-1}\Lambda H_2^{-1}H_3H_0^{-1}$$

$$D_d = \sigma^2 (\Lambda + kdI)(\Lambda + kI)^{-1}\Lambda^{-1}(\Lambda + kdI)(\Lambda + kI)^{-1} - \sigma^2 (\Lambda + kdI)^{-1}(\Lambda - kdI)(\Lambda + kI)^{-1}$$

$$\Lambda(\Lambda + kdI)^{-1}(\Lambda - kdI)(\Lambda + kI)^{-1}$$

$$D_d = \sigma^2 \text{diag} \left\{ \frac{(\lambda_i + kd)^2}{\lambda(\lambda_i + k)^2} - \frac{\lambda_i(\lambda_i - kd)^2}{(\lambda_i + kd)^2(\lambda_i + k)^2} \right\}_{i=1}^p \tag{55}$$

will be pdf if and only if $(\lambda_i + kd)^4 - \lambda_i^2(\lambda_i - kd)^2 > 0$. For $0 < d < 1$ and $k > 0$, it was observed that $(\lambda_i + kd)^4 - \lambda_i^2(\lambda_i - kd)^2 > 0$. By lemma 3, the proof is completed.

2.5 Determination of Biasing Parameters k and d

Finding the appropriate ridge and Liu parameters, k and d , respectively, is a critical issue in the study of the ridge and Liu regression. These parameters may either be non-stochastic or may depend on the observed data. The choice of values for these ridge parameters has been one of the most difficult problems confronting the study of the generalized ridge regression, [35]. The biasing parameters k proposed by [12] in equation (6) will be used to determine and evaluate the performance of the proposed estimators compared to the OLS estimator and other estimators. It is given as:

$$k = \frac{\hat{\sigma}^2}{2\hat{\alpha}_i^2 + (\hat{\sigma}^2/\lambda_i)} \tag{56}$$

In determining the optimal value of d for $\hat{\alpha}_{NTPKL}(k, d)$, k is fixed. The optimal value of d can be regarded as to be that d that minimizes

$$MSEM[\hat{\alpha}_{NTPKL}(k, d)] = \sigma^2 H_2^{-1}H_3H_0^{-1}\Lambda H_2^{-1}H_3H_0^{-1} + (\Lambda H_2^{-1}H_3H_0^{-1} - I)\alpha\alpha'(\Lambda H_2^{-1}H_3H_0^{-1} - I)'$$

$$f(k, d) = MSEM[\hat{\alpha}_{NTPKL}(k, d)] = \text{tr}[MSEM(\hat{\alpha}_{NTPKL}(k, d))]$$

$$f(k, d) = \sigma^2 \sum_i^p \frac{\lambda_i(\lambda_i - kd)^2}{(\lambda_i + kd)^2(\lambda_i + k)^2} + \sum_i^p \alpha_i^2 \frac{(2\lambda_i kd + \lambda_i k + k^2 d)^2}{(\lambda_i + kd)^2(\lambda_i + k)^2} \tag{57}$$

Differentiating $f(k, d)$ with respect to d and equate to zero gives

$$d = \sum_i^p \left[\frac{\lambda_i(\sigma^2 - \alpha^2 k)}{2\alpha^2 \lambda k + \alpha^2 k^2 + \sigma^2 k} \right] \tag{58}$$

For practical purposes, σ^2 and α_i^2 are replaced with $\hat{\sigma}^2$ and $\hat{\alpha}_i^2$, respectively. Consequently, (58) becomes

$$\hat{d} = \sum_i^p \left[\frac{\lambda_i (\hat{\sigma}^2 - \hat{\alpha}_i^2 k)}{2\hat{\alpha}_i^2 \lambda k + \hat{\alpha}_i^2 k^2 + \hat{\sigma}^2 k} \right] \quad (59)$$

3 Simulation Study

To support the theoretical comparison of the estimators, a Monte Carlo simulation study was conducted using R 4.0.2 to examine the performance of these estimators. Also, in this section, the results of the simulation will be discussed.

3.1 Simulation Technique

The simulation procedure used by [22], [25], [35] and [36], were used to generate the explanatory variables in this study: This is given as:

$$x_{ij} = (1 - \rho^2)^{1/2} z_{ij} + \rho z_{i,p+1}, \quad i = 1, 2, \dots, n, \\ j = 1, 2, \dots, p. \quad (60)$$

where z_{ij} is an independent standard normal pseudo-random number, ρ is the correlation between any two explanatory variables considered to be 0.8, 0.9, 0.95, and 0.99. Also, p is the number of explanatory variables, and $p=3$ is considered for the simulation study. The variables are standardized so that $X'X$ and $X'y$ are in correlation forms. The values of $\beta'\beta = 1$, [37]. The standard deviations in this simulation study were $\sigma = 1, 3, 5,$ and 10 . In comparing the estimators, k was chosen to be 0.3, 0.6, and 0.9, which lies between 0 and 1, following [35] and [38]. We also chose $d = 0.2, 0.5, 0.8$ for sample sizes 50 and 100. The replication for the study is 1000 times. The mean square error was obtained using:

$$MSE(\hat{\alpha}) = \frac{1}{1000} \sum_{j=1}^{1000} (\hat{\alpha}_{ij} - \alpha_i)' (\hat{\alpha}_{ij} - \alpha_i) \quad (61)$$

Table 1 (Appendix) shows that when $k=0.3$ and $d=0.2$, $\hat{\alpha}_{LE}(d)$ and $\hat{\alpha}_{MTPKL1}(k,d)$ consistently perform better than $\hat{\alpha}_{NTPKL}(k,d)$ and other existing estimators across the four levels of sigma (1, 3, 5, and 10) and at the two sample sizes ($n = 50$ and 100), particularly when rho is 0.8, 0.9, 0.95. At sample size 50, $\hat{\alpha}_{KL}(k)$ and $\hat{\alpha}_{MTPKL1}(k,d)$ dominates other estimators, while at $n= 100$, the proposed estimators,

$\hat{\alpha}_{MTPKL1}(k,d)$ and $\hat{\alpha}_{MTPKL2}(k,d)$ perform better than other estimators when rho is 0.99.

From Table 2, Table 3, Table 9 and Table 12 in Appendix (when $k=0.3$ and $d=0.5$, $k=0.3$ and $d=0.8$, $k=0.7$ and $d=0.8$, $k=0.9$ and $d=0.8$), the simulation results show that two of the proposed estimators; $\hat{\alpha}_{MTPKL1}(k,d)$ and $\hat{\alpha}_{MTPKL2}(k,d)$ consistently outperform $\hat{\alpha}_{NTPKL}(k,d)$ and other existing estimators in this study at sample size 50 and 100 across the four levels of sigma (1, 3, 5 and 10) and at rho is 0.8, 0.9, 0.95, and 0.99

From Table 4 and Table 7 in Appendix (when $k=0.6$ and $d=0.2$, $k=0.7$ and $d=0.2$), the simulation results show that two of the proposed estimators; $\hat{\alpha}_{MTPKL1}(k,d)$ and $\hat{\alpha}_{MTPKL2}(k,d)$ consistently dominate $\hat{\alpha}_{NTPKL}(k,d)$ and other existing estimators in this study except for when rho is 0.99, where $\hat{\alpha}_{MTPKL1}(k,d)$ and $\hat{\alpha}_{KL}(k)$ dominates $\hat{\alpha}_{NTPKL}(k,d)$ and other existing estimators at sample size 50 and 100.

From Table 5 and Table 8 in Appendix (when $k=0.6$ and $d=0.5$, $k=0.7$ and $d=0.5$), the simulation results show that two of the proposed estimators; $\hat{\alpha}_{MTPKL1}(k,d)$ and $\hat{\alpha}_{MTPKL2}(k,d)$ consistently dominate $\hat{\alpha}_{NTPKL}(k,d)$ and other existing estimators in this study except for when rho is 0.99, where $\hat{\alpha}_{MTPKL1}(k,d)$ and $\hat{\alpha}_{KL}(k)$ dominates $\hat{\alpha}_{NTPKL}(k,d)$ and other existing estimators at sample size 50 only.

From Table 6 in Appendix (when $k=0.6$ and $d=0.8$), the simulation results show that two of the proposed estimators; $\hat{\alpha}_{MTPKL1}(k,d)$ and $\hat{\alpha}_{MTPKL2}(k,d)$ consistently dominate $\hat{\alpha}_{NTPKL}(k,d)$ and other existing estimators in this study except for when rho is 0.99, sigma = 10 and $n=50$ where $\hat{\alpha}_{MTPKL1}(k,d)$ and $\hat{\alpha}_{NTPKL}(k,d)$ dominates $\hat{\alpha}_{MTPKL2}(k,d)$ and other existing estimators.

From Table 10 in Appendix (when $k=0.9$ and $d=0.2$), the simulation results show that two of the proposed estimators; $\hat{\alpha}_{MTPKL1}(k,d)$ and $\hat{\alpha}_{MTPKL2}(k,d)$ consistently dominate $\hat{\alpha}_{NTPKL}(k,d)$ and other existing estimators in this study except for when rho is 0.95 and 0.99, where $\hat{\alpha}_{MTPKL1}(k,d)$ and $\hat{\alpha}_{KL}(k)$ dominates $\hat{\alpha}_{MTPKL2}(k,d)$ and other existing

estimators at sample size 50. However, when the sample size is 100, $\hat{\alpha}_{MTPKL1}(k, d)$ and $\hat{\alpha}_{MTPKL2}(k, d)$ consistently dominates $\hat{\alpha}_{NTPKL}(k, d)$ and other existing estimators except for when $\rho = 0.99$, where $\hat{\alpha}_{MTPKL1}(k, d)$ and $\hat{\alpha}_{KL}(k)$ dominates $\hat{\alpha}_{MTPKL2}(k, d)$ and other existing estimators.

From Table 11 in Appendix (when $k=0.9$ and $d=0.5$), the simulation results show that two of the proposed estimators; $\hat{\alpha}_{MTPKL1}(k, d)$ and $\hat{\alpha}_{MTPKL2}(k, d)$ consistently dominate $\hat{\alpha}_{NTPKL}(k, d)$ and other existing estimators in this study while $\hat{\alpha}_{MTPKL1}(k, d)$ has the least MSE value among the three proposed estimators across the sample sizes, σ and ρ levels used in this study.

4 Numerical Example

In this section, Longley data was used to demonstrate the performance of the proposed estimator. Longley data were used by [39] and [40]. The regression model for these data is defined as:

$$y = \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_6 X_6 + \varepsilon \quad (62)$$

For more details on the data set, [39]. The variance inflation factors are $VIF_1 = 135.53$, $VIF_2 = 1788.51$, $VIF_3 = 33.62$, $VIF_4 = 3.59$, $VIF_5 = 399.15$ and $VIF_6 = 758.98$. Eigenvalues of XX matrix are $\lambda_1 = 2.76779 \times 10^{12}$, $\lambda_2 = 7,039,139,179$, $\lambda_3 = 11,608,993.96$, $\lambda_4 = 2,504,761.021$, $\lambda_5 = 1738.356$, $\lambda_6 = 13.309$ and the condition number of XX is 456,070. The VIFs, the eigenvalues, and the condition number all indicate that severe multicollinearity exists. The estimated parameters and the MSE values of the estimators are presented in Table 13 (Appendix). Two of the proposed estimators perform best among other estimators as they given the smallest MSE value

Though, $\hat{\alpha}_{MTPKL1}(k, d)$ and $\hat{\alpha}_{MTPKL2}(k, d)$ which are special cases of $\hat{\alpha}_{NTPKL}(k, d)$ perform better than $\hat{\alpha}_{NTPKL}(k, d)$ and other existing estimators, as they give smaller MSE values compared with all the existing estimators considered in this study, $\hat{\alpha}_{MTPKL1}(k, d)$ performs best in all. This is consistent with the results of the simulation study.

5 Summary and Concluding Remarks

This paper proposes new two-parameter estimators to solve the multicollinearity problem for linear regression models. The proposed estimators were theoretically compared with six other existing estimators. A simulation study was then conducted to compare the performance of the proposed estimators; $\hat{\alpha}_{MTPKL1}(k, d)$, $\hat{\alpha}_{MTPKL2}(k, d)$ and $\hat{\alpha}_{NTPKL}(k, d)$ with six existing estimators. From the theoretical comparison, simulation study, and application of life data, the proposed estimators, especially the two special cases of $\hat{\alpha}_{NTPKL}(k, d)$ which are $\hat{\alpha}_{MTPKL1}(k, d)$ and $\hat{\alpha}_{MTPKL2}(k, d)$ give better results in terms of MSE. Hopefully, this paper will be helpful to researchers in different fields. The proposed estimators are as a result of this, recommended for use by researchers in various fields.

References:

- [1] B. F. Swindel, "Good ridge estimators based on prior information," *Communications in Statistics—Theory and Methods*, vol. 5, no. (11), p. 1065–75, 1976.
- [2] M. R. Baye, and D. F. Parker., "Combining ridge and principal component regression: a money demand illustration," *Communications in Statistics - Theory and Methods*, vol. 13(2), p. 197–205, 1984.
- [3] J. P. Buonaccorsi, "A Modified estimating equation approach to correcting for measurement error in regression," *Biometrika*, vol. 83, p. 433–440, 1996.
- [4] A. Koutsoyannis, *Theory of Econometrics*, 2nd ed., Palgrave Publishers, 2003.
- [5] G. S. Maddala, *Introduction to Econometrics*, Singapore: John Wiley and Sons, 2005.
- [6] C. Stein, "Inadmissibility of the usual estimator for mean of multivariate normal distribution," in *Proceedings of the Oird Berkley Symposium on Mathematical and Statistics Probability*, vol. 1, J. Neyman, Ed., Berlin: Springer, 1956.
- [7] W. F. Massy, "Principal components regression in exploratory statistical research," *Journal of the American Statistical Association*, vol. 60, no. 309, p. 234–256, 1965.

- [8] Hoerl, A. E. and Kennard, R. W., "Ridge regression: biased estimation for nonorthogonal problems," *Technometrics*, vol. 12, no. 1, p. 55–67, 1970.
- [9] B. F. Swindel, "Good ridge estimators based on prior information," *Communications in Statistics-Theory and Methods*, vol. 5, no. 11, p. 1065–1075, 1976.
- [10] Mayer, L. S. and Willke, T.A., "On biased estimation in linear models," *Technometrics*, vol. 15, no. 3, p. 497–508, 1973.
- [11] K. Liu, "A new class of biased estimate in linear regression.," *Communications in Statistics*, vol. 22, no. 2, pp. 393-402, 1993.
- [12] Kibria, B. M. G. and Lukman, A. F., "A New Ridge-Type Estimator for the Linear Regression Model: Simulations and Applications," *Hindawi Scientifica*, p. 16 pages, 2020.
- [13] Ozkale, M. R. and Kaçiranlar, S., "The restricted and unrestricted two-parameter estimators," *Communications in Statistics-Theory and Methods*, vol. 36, no. 15, p. 2707–2725, 2007.
- [14] Yang, H. and Chang, X., "A new two-parameter estimator in linear regression," *Communications in Statistics-Theory and Methods*, vol. 39, no. 6, p. 923–934, 2010.
- [15] A. V. Dorugade, "Modified two parameter estimator in linear regression," *Statistics in Transition New Series*, vol. 15, no. 1, p. 23-36, 2014.
- [16] M. Roozbeh, "Optimal QR-based estimation in partially linear regression models with correlated errors using GCV criterion," *Computational Statistics & Data Analysis*, vol. 117, p. 45–61, 2018.
- [17] Lukman, A. F.; Ayinde, K.; Binuomote, S. and Clement, O. A., "Modified ridge-type estimator to combat multicollinearity: application to chemical data," *Journal of Chemometrics*, vol. 33, no. 5, p. e3125, 2019.
- [18] Dawoud, I and Kibria, B. M. G., "A New Biased Estimator to Combat the Multicollinearity of the Gaussian Linear Regression Model," *Stats*, vol. 3, p. 526–541, 2020.
- [19] Owolabi, A. T., Ayinde, K., Idowu, J. I., Oladapo, O. J., & Lukman, A. F., "A New Two-Parameter Estimator in the Linear Regression Model with Correlated Regressors," *Journal of Statistics Applications & Probability*, vol. 11, pp. 499-512, 2022.
- [20] Idowu, J. I., Oladapo, O. J., Owolabi, A. T. and Ayinde, K., "On the Biased Two-Parameter Estimator to Combat Multicollinearity in Linear Regression Model," *African Scientific Reports*, vol. 1, p.188-204, 2022.
- [21] Hoerl, A. E.; Kannard, R. W. and Baldwin, K. F., "Ridge regression: some simulations," *Communications in Statistics*, vol. 4, no. 2, p. 105–123, 1975.
- [22] McDonald, G. C. and Galarneau, D. I., "A Monte Carlo evaluation of some ridge-type estimators," *Journal of the American Statistical Association*, vol. 70, no. 350, p. 407–416, 1975.
- [23] Lawless, J. F. and Wang, P., "A simulation study of ridge and other regression estimators," *Communications in Statistics-Theory and Methods*, vol. 5, no. 4, p. 307–323, 1976.
- [24] L. L. Firinguetti, "A simulation study of ridge regression estimators with autocorrelated errors," *Communications in Statistics Simulation and Computation*, vol. 18, no. 2, p. 673–702, 1989.
- [25] B. M. Kibria, "Performance of some new ridge regression estimators," *Communications in Statistics-Simulation and Computation*, vol. 32, no. 1, p. 419–435, 2003.
- [26] Batach, F. S., Ramnathan, T. and Gore, S. D., "The Efficiency of Modified Jackknife and Ridge Type Regression Estimators: A Comparison," *Surveys in Mathematics and its Applications*, vol. 24, no. 2, p. 157–174, 2008.
- [27] A. V. Dorugade, "New Ridge Parameters for Ridge Regression," *Journal of the Association of Arab Universities for Basic and Applied Sciences*, pp. 1-6, 2016.
- [28] Oladapo, O.J., Owolabi, A.T., Idowu, J.I. and Ayinde, K. , "A New Modified Liu Ridge-Type Estimator for the Linear Regression Model: Simulation and Application," *Int J Clin Biostat Biom*, vol. 8, no. 048, 2022.
- [29] Idowu, J.I., Oladapo, O. J., Owolabi, A. T., Ayinde, K. and Akinmoju, O., "Combating Multicollinearity: A new two-parameter Approach," *Journal of Quantitative Sciences*, vol. 5, no. 1, 2023.
- [30] Aslam, M. & Ahmad, S., "The modified Liu-

ridge-type estimator: a new class of biased estimators to address multicollinearity," *Communications in Statistics - Simulation and Computation*, 2020.

- [31] Sakallıoğlu, S. and Kaçiranlar, S., "A new biased estimator, based on ridge estimation," *Stat. Pap.*, vol. 49, no. 4, p. 669–689, 2008.
- [32] R. Farebrother, "Further results on the mean square error of ridge regression," *J. R. Stat. Soc. B*, vol. 38, 1976.
- [33] Trenkler, G. & Toutenburg, H., "Mean squared error matrix comparisons between biased estimators-an overview of recent results," *Stat Pap.*, vol. 31, pp. 165-179, 1990.
- [34] A. Dorugade, "Efficient Two-Parameter Estimator In Linear Regression Model," *Statistics in transition new series*, vol. 20, no. 2, p. 173–185, 2019.
- [35] Wichern, D. W. and Churchill, G. A., "A comparison of ridge estimators," *Technometrics*, vol. 20, no. 3, p. 301–311, 1978.
- [36] D. G. Gibbons, "A simulation study of some ridge estimators.," *Journal of the American Statistical Association*, vol. 76, pp. 131-139, 1981.
- [37] Newhouse, J. P. and Oman, S. D., "An evaluation of ridge estimators. A report prepared for United States Air Force project RAND", 1971.
- [38] Kan, B.; Alpu, O.; and Yazici, B., "Robust ridge and robust Liu estimator for regression based on LTS estimator," *J. Appl. Stat.*, vol. 40, p. 644-655, 2013.
- [39] J. Longley, "An appraisal of least squares programs for electronic computer from the point of view of the user," *J. Am. Stat. Assoc.*, vol. 62, p. 819–841, 1967.
- [40] A. Yasin and E. Murat, "Influence Diagnostics in Two-Parameter Ridge Regression," *J. Data Sci.*, vol. 14, p. 33–52, 2016.

Contribution of Individual Authors to the Creation of a Scientific Article (Ghostwriting Policy)

Idowu, J.I. Ayinde, K., Oladapo, O. J. and Owolabi, A.T. conceived and designed the study. Oladapo, O. J., Alao, N. A., Idowu, J.I. and Oshuporu, O. A. conducted the data analysis. Ayinde K. supervised the study. Idowu, J. I., Oladapo, O. J., Owolabi, A.T., Ayinde K., and Alao, N. A. interpreted the study results. Owolabi, A. T., Oladapo, O.J., Idowu, J. I. and Oshuporu, O. A. wrote the first draft of the manuscript. Ayinde K. and Owolabi, A. T. reviewed and corrected the draft manuscript. All authors read through and approved the final manuscript.

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Conflicts of Interest

The authors declare that they have no conflicts of interest.

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APPENDIX

Table 1. Estimated MSE when k=0.3, d=0.2

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1311	0.1234	0.1261	0.1301	0.1322	0.1291	0.1251	0.1252
		0.9	0.2462	0.2288	0.2046	0.2120	0.2255	0.2322	0.2220	0.2089	0.2094
		0.95	0.4682	0.4064	0.3333	0.3492	0.3956	0.4184	0.3833	0.3393	0.3429
		0.99	2.2417	1.2173	0.7223	0.5101	1.0998	1.3968	0.9166	0.4469	0.5771
	3	0.8	1.2270	1.1802	1.1098	1.1344	1.1712	1.1895	1.1618	1.1255	1.1264
		0.9	2.2158	2.0587	1.8412	1.9076	2.0293	2.0896	1.9980	1.8794	1.8847
		0.95	4.2136	3.6579	2.9999	3.1428	3.5604	3.7658	3.4501	3.0533	3.0857
		0.99	20.1751	10.9561	6.5009	4.5913	9.8985	12.5716	8.2494	4.0221	5.1935
	5	0.8	3.4082	3.2782	3.0825	3.1509	3.2532	3.3040	3.2272	3.1264	3.1288
		0.9	6.1551	5.7185	5.1142	5.2987	5.6370	5.8045	5.5499	5.2205	5.2351
		0.95	11.7045	10.1609	8.3330	8.7300	9.8901	10.4606	9.5836	8.4815	8.5716
		0.99	56.0420	30.4336	18.0581	12.7536	27.4958	34.9210	22.9151	11.1726	14.4263
10	0.8	13.6328	13.1127	12.3294	12.6032	13.0124	13.2158	12.9084	12.5051	12.5148	
	0.9	24.6204	22.8740	20.4565	21.1943	22.5479	23.2179	22.1993	20.8816	20.9400	
	0.95	46.8180	40.6437	33.3323	34.9200	39.5605	41.8426	38.3346	33.9260	34.2864	
	0.99	224.1680	121.7344	72.2327	51.0147	109.9835	139.6840	91.6605	44.6905	57.7054	
100	1	0.8	0.0622	0.0611	0.0594	0.0600	0.0609	0.0613	0.0607	0.0598	0.0598
		0.9	0.1131	0.1093	0.1036	0.1056	0.1086	0.1101	0.1078	0.1049	0.1050
		0.95	0.2163	0.2023	0.1828	0.1889	0.1997	0.2051	0.1970	0.1864	0.1868
		0.99	1.0445	0.7667	0.5345	0.5329	0.7248	0.8187	0.6710	0.4995	0.5275
	3	0.8	0.5596	0.5499	0.5345	0.5403	0.5480	0.5518	0.5461	0.5384	0.5385
		0.9	1.0178	0.9839	0.9322	0.9506	0.9774	0.9907	0.9706	0.9442	0.9448
		0.95	1.9464	1.8211	1.6447	1.7002	1.7976	1.8458	1.7726	1.6775	1.6813
		0.99	9.4002	6.9000	4.8107	4.7956	6.5235	7.3684	6.0393	4.4950	4.7478
	5	0.8	1.5543	1.5275	1.4848	1.5009	1.5222	1.5328	1.5168	1.4957	1.4959
		0.9	2.8273	2.7332	2.5896	2.6407	2.7149	2.7519	2.6961	2.6228	2.6243
		0.95	5.4066	5.0586	4.5687	4.7227	4.9932	5.1273	4.9238	4.6598	4.6704
		0.99	26.1116	19.1667	13.3631	13.3212	18.1207	20.4678	16.7757	12.4861	13.1882
10	0.8	6.2173	6.1100	5.9393	6.0036	6.0889	6.1314	6.0674	5.9827	5.9837	
	0.9	11.3093	10.9326	10.3582	10.5627	10.8597	11.0074	10.7844	10.4911	10.4973	
	0.95	21.6263	20.2345	18.2747	18.8909	19.9729	20.5090	19.6951	18.6390	18.6815	
	0.99	104.4463	76.6669	53.4523	53.2847	72.4828	81.8710	67.1027	49.9442	52.7527	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 2. Estimated MSE when k=0.3, d=0.5

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1311	0.1281	0.1261	0.1287	0.1337	0.1261	0.1237	0.1237
		0.9	0.2462	0.2288	0.2197	0.2120	0.2207	0.2374	0.2123	0.2043	0.2048
		0.95	0.4682	0.4064	0.3811	0.3492	0.3802	0.4367	0.3512	0.3252	0.3278
		0.99	2.2417	1.2173	1.1925	0.5101	0.9535	1.6899	0.5978	0.3726	0.4549
	3	0.8	1.2270	1.1802	1.1530	1.1344	1.1578	1.2034	1.1348	1.1125	1.1132
		0.9	2.2158	2.0587	1.9774	1.9076	1.9865	2.1365	1.9104	1.8384	1.8424
		0.95	4.2136	3.6579	3.4302	3.1428	3.4215	3.9307	3.1608	2.9264	2.9506
		0.99	20.1751	10.9561	10.7325	4.5913	8.5816	15.2088	5.3805	3.3531	4.0942
	5	0.8	3.4082	3.2782	3.2026	3.1509	3.2161	3.3429	3.1522	3.0902	3.0920
		0.9	6.1551	5.7185	5.4928	5.2987	5.5181	5.9347	5.3065	5.1066	5.1177
		0.95	11.7045	10.1609	9.5284	8.7300	9.5043	10.9186	8.7801	8.1288	8.1961
		0.99	56.0420	30.4336	29.8126	12.7536	23.8378	42.2467	14.9459	9.3141	11.3728
10	0.8	13.6328	13.1127	12.8101	12.6032	12.8643	13.3714	12.6084	12.3602	12.3677	
	0.9	24.6204	22.8740	21.9710	21.1943	22.0722	23.7388	21.2258	20.4259	20.4704	
	0.95	46.8180	40.6437	38.1136	34.9200	38.0173	43.6745	35.1204	32.5154	32.7845	
	0.99	224.1680	121.7344	119.2504	51.0147	95.3513	168.9870	59.7836	37.2565	45.4912	
100	1	0.8	0.0622	0.0611	0.0604	0.0600	0.0606	0.0616	0.0600	0.0595	0.0595
		0.9	0.1131	0.1093	0.1071	0.1056	0.1075	0.1112	0.1057	0.1039	0.1039
		0.95	0.2163	0.2023	0.1950	0.1889	0.1959	0.2092	0.1891	0.1827	0.1830
		0.99	1.0445	0.7667	0.7055	0.5329	0.6683	0.9001	0.5497	0.4551	0.4749
	3	0.8	0.5596	0.5499	0.5438	0.5403	0.5452	0.5547	0.5404	0.5356	0.5357
		0.9	1.0178	0.9839	0.9639	0.9506	0.9677	1.0008	0.9509	0.9347	0.9351
		0.95	1.9464	1.8211	1.7548	1.7002	1.7631	1.8832	1.7022	1.6444	1.6473
		0.99	9.4002	6.9000	6.3497	4.7956	6.0150	8.1006	4.9469	4.0958	4.2742
	5	0.8	1.5543	1.5275	1.5107	1.5009	1.5144	1.5409	1.5010	1.4879	1.4881
		0.9	2.8273	2.7332	2.6774	2.6407	2.6879	2.7800	2.6415	2.5963	2.5975
		0.95	5.4066	5.0586	4.8743	4.7227	4.8976	5.2311	4.7284	4.5677	4.5758
		0.99	26.1116	19.1667	17.6381	13.3212	16.7084	22.5017	13.7414	11.3772	11.8727
10	0.8	6.2173	6.1100	6.0427	6.0036	6.0574	6.1635	6.0041	5.9516	5.9523	
	0.9	11.3093	10.9326	10.7097	10.5627	10.7517	11.1201	10.5660	10.3851	10.3899	
	0.95	21.6263	20.2345	19.4972	18.8909	19.5904	20.9244	18.9138	18.2708	18.3033	
	0.99	104.4463	76.6669	70.5524	53.2847	66.8336	90.0070	54.9657	45.5089	47.4907	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 3. Estimated MSE when k=0.3, d=0.8

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1311	0.1330	0.1261	0.1272	0.1353	0.1232	0.1223	0.1223
		0.9	0.2462	0.2288	0.2354	0.2120	0.2162	0.2427	0.2030	0.1999	0.2001
		0.95	0.4682	0.4064	0.4323	0.3492	0.3657	0.4555	0.3218	0.3119	0.3132
		0.99	2.2417	1.2173	1.7822	0.5101	0.8349	2.0114	0.3858	0.3158	0.3479
	3	0.8	1.2270	1.1802	1.1971	1.1344	1.1448	1.2175	1.1085	1.0997	1.1000
		0.9	2.2158	2.0587	2.1188	1.9076	1.9451	2.1839	1.8268	1.7988	1.8007
		0.95	4.2136	3.6579	3.8903	3.1428	3.2909	4.0992	2.8962	2.8074	2.8185
		0.99	20.1751	10.9561	16.0396	4.5913	7.5144	18.1030	3.4725	2.8421	3.1315
	5	0.8	3.4082	3.2782	3.3251	3.1509	3.1798	3.3820	3.0790	3.0546	3.0555
		0.9	6.1551	5.7185	5.8855	5.2987	5.4031	6.0665	5.0742	4.9965	5.0017
		0.95	11.7045	10.1609	10.8064	8.7300	9.1413	11.3868	8.0450	7.7985	7.8293
		0.99	56.0420	30.4336	44.5544	12.7536	20.8735	50.2860	9.6458	7.8948	8.6985
10	0.8	13.6328	13.1127	13.3005	12.6032	12.7188	13.5279	12.3156	12.2179	12.2214	
	0.9	24.6204	22.8740	23.5419	21.1943	21.6120	24.2658	20.2964	19.9855	20.0064	
	0.95	46.8180	40.6437	43.2258	34.9200	36.5654	45.5471	32.1800	31.1940	31.3174	
	0.99	224.1680	121.7344	178.2178	51.0147	83.4941	201.1442	38.5833	31.5795	34.7943	
100	1	0.8	0.0622	0.0611	0.0615	0.0600	0.0603	0.0620	0.0594	0.0592	0.0592
		0.9	0.1131	0.1093	0.1107	0.1056	0.1065	0.1123	0.1035	0.1028	0.1028
		0.95	0.2163	0.2023	0.2076	0.1889	0.1922	0.2134	0.1817	0.1792	0.1793
		0.99	1.0445	0.7667	0.9008	0.5329	0.6183	0.9854	0.4500	0.4166	0.4252
	3	0.8	0.5596	0.5499	0.5532	0.5403	0.5424	0.5576	0.5347	0.5329	0.5329
		0.9	1.0178	0.9839	0.9961	0.9506	0.9581	1.0110	0.9317	0.9253	0.9255
		0.95	1.9464	1.8211	1.8685	1.7002	1.7297	1.9210	1.6348	1.6123	1.6136
		0.99	9.4002	6.9000	8.1072	4.7956	5.5649	8.8685	4.0503	3.7489	3.8264
	5	0.8	1.5543	1.5275	1.5368	1.5009	1.5066	1.5489	1.4854	1.4802	1.4803
		0.9	2.8273	2.7332	2.7668	2.6407	2.6614	2.8084	2.5881	2.5702	2.5708
		0.95	5.4066	5.0586	5.1902	4.7227	4.8048	5.3360	4.5411	4.4785	4.4823
		0.99	26.1116	19.1667	22.5199	13.3212	15.4581	24.6347	11.2509	10.4136	10.6289
10	0.8	6.2173	6.1100	6.1472	6.0036	6.0262	6.1958	5.9416	5.9207	5.9210	
	0.9	11.3093	10.9326	11.0674	10.5627	10.6454	11.2334	10.3522	10.2809	10.2831	
	0.95	21.6263	20.2345	20.7609	18.8909	19.2192	21.3441	18.1645	17.9139	17.9292	
	0.99	104.4463	76.6669	90.0797	53.2847	61.8325	98.5387	45.0034	41.6544	42.5154	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 4. Estimated MSE when k=0.6, d=0.2

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1263	0.1234	0.1167	0.1244	0.1283	0.1224	0.1149	0.1153
		0.9	0.2462	0.2132	0.2046	0.1827	0.2074	0.2196	0.2009	0.1774	0.1794
		0.95	0.4682	0.3564	0.3333	0.2604	0.3391	0.3775	0.3173	0.2461	0.2566
		0.99	2.2417	0.7679	0.7223	0.0801	0.6563	0.9998	0.4381	0.0644	0.2020
	3	0.8	1.2270	1.1362	1.1098	1.0491	1.1193	1.1540	1.1012	1.0330	1.0362
		0.9	2.2158	1.9183	1.8412	1.6431	1.8663	1.9760	1.8072	1.5955	1.6133
		0.95	4.2136	3.2080	2.9999	2.3436	3.0516	3.3978	2.8556	2.2146	2.3094
		0.99	20.1751	6.9114	6.5009	0.7210	5.9070	8.9983	3.9425	0.5794	1.8174
	5	0.8	3.4082	3.1559	3.0825	2.9139	3.1090	3.2055	3.0587	2.8690	2.8779
		0.9	6.1551	5.3284	5.1142	4.5638	5.1839	5.4888	5.0199	4.4316	4.4810
		0.95	11.7045	8.9111	8.3330	6.5101	8.4768	9.4384	7.9322	6.1518	6.4149
		0.99	56.0420	19.1985	18.0581	2.0029	16.4084	24.9953	10.9513	1.6093	5.0484
10	0.8	13.6328	12.6233	12.3294	11.6547	12.4356	12.8219	12.2341	11.4752	11.5108	
	0.9	24.6204	21.3135	20.4565	18.2547	20.7355	21.9550	20.0791	17.7258	17.9234	
	0.95	46.8180	35.6446	33.3323	26.0407	33.9073	37.7537	31.7291	24.6072	25.6600	
	0.99	224.1680	76.7940	72.2327	8.0117	65.6337	99.9814	43.8053	6.4375	20.1937	
100	1	0.8	0.0622	0.0601	0.0594	0.0580	0.0597	0.0605	0.0592	0.0576	0.0576
		0.9	0.1131	0.1058	0.1036	0.0987	0.1044	0.1072	0.1029	0.0974	0.0976
		0.95	0.2163	0.1898	0.1828	0.1651	0.1851	0.1949	0.1798	0.1608	0.1622
		0.99	1.0445	0.5881	0.5345	0.2651	0.5341	0.6687	0.4515	0.2342	0.2920
	3	0.8	0.5596	0.5405	0.5345	0.5218	0.5368	0.5443	0.5330	0.5182	0.5185
		0.9	1.0178	0.9518	0.9322	0.8881	0.9394	0.9648	0.9262	0.8762	0.8782
		0.95	1.9464	1.7080	1.6447	1.4859	1.6658	1.7544	1.6184	1.4469	1.4600
		0.99	9.4002	5.2929	4.8107	2.3853	4.8065	6.0184	4.0635	2.1077	2.6277
	5	0.8	1.5543	1.5014	1.4848	1.4495	1.4912	1.5119	1.4806	1.4394	1.4403
		0.9	2.8273	2.6439	2.5896	2.4669	2.6094	2.6800	2.5728	2.4338	2.4395
		0.95	5.4066	4.7444	4.5687	4.1273	4.6272	4.8733	4.4956	4.0191	4.0555
		0.99	26.1116	14.7025	13.3631	6.6258	13.3514	16.7177	11.2874	5.8547	7.2991
10	0.8	6.2173	6.0056	5.9393	5.7978	5.9646	6.0476	5.9223	5.7577	5.7613	
	0.9	11.3093	10.5755	10.3582	9.8677	10.4378	10.7202	10.2913	9.7352	9.7580	
	0.95	21.6263	18.9777	18.2747	16.5093	18.5089	19.4930	17.9825	16.0762	16.2220	
	0.99	104.4463	58.8099	53.4523	26.5030	53.4054	66.8709	45.1496	23.4186	29.1961	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 5. Estimated MSE when k=0.6, d=0.5

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1263	0.1281	0.1167	0.1217	0.1312	0.1169	0.1123	0.1126
		0.9	0.2462	0.2132	0.2197	0.1827	0.1992	0.2294	0.1838	0.1699	0.1714
		0.95	0.4682	0.3564	0.3811	0.2604	0.3154	0.4103	0.2666	0.2268	0.2343
		0.99	2.2417	0.7679	1.1925	0.0801	0.5301	1.4066	0.1837	0.0488	0.1193
	3	0.8	1.2270	1.1362	1.1530	1.0491	1.0947	1.1811	1.0508	1.0095	1.0120
		0.9	2.2158	1.9183	1.9774	1.6431	1.7923	2.0643	1.6530	1.5281	1.5414
		0.95	4.2136	3.2080	3.4302	2.3436	2.8383	3.6931	2.3995	2.0409	2.1086
		0.99	20.1751	6.9114	10.7325	0.7210	4.7705	12.6591	1.6527	0.4386	1.0732
	5	0.8	3.4082	3.1559	3.2026	2.9139	3.0407	3.2808	2.9187	2.8038	2.8106
		0.9	6.1551	5.3284	5.4928	4.5638	4.9784	5.7340	4.5914	4.2443	4.2812
		0.95	11.7045	8.9111	9.5284	6.5101	7.8841	10.2588	6.6653	5.6693	5.8573
		0.99	56.0420	19.1985	29.8126	2.0029	13.2514	35.1642	4.5908	1.2183	2.9811
10	0.8	13.6328	12.6233	12.8101	11.6547	12.1622	13.1229	11.6739	11.2140	11.2411	
	0.9	24.6204	21.3135	21.9710	18.2547	19.9132	22.9359	18.3648	16.9765	17.1238	
	0.95	46.8180	35.6446	38.1136	26.0407	31.5367	41.0351	26.6614	22.6773	23.4292	
	0.99	224.1680	76.7940	119.2504	8.0117	53.0058	140.6570	18.3634	4.8734	11.9245	
100	1	0.8	0.0622	0.0601	0.0604	0.0580	0.0591	0.0611	0.0580	0.0570	0.0571
		0.9	0.1131	0.1058	0.1071	0.0987	0.1024	0.1094	0.0988	0.0955	0.0956
		0.95	0.2163	0.1898	0.1950	0.1651	0.1784	0.2028	0.1659	0.1546	0.1557
		0.99	1.0445	0.5881	0.7055	0.2651	0.4662	0.7996	0.3035	0.1974	0.2345
	3	0.8	0.5596	0.5405	0.5438	0.5218	0.5314	0.5500	0.5220	0.5129	0.5131
		0.9	1.0178	0.9518	0.9639	0.8881	0.9213	0.9845	0.8892	0.8588	0.8604
		0.95	1.9464	1.7080	1.7548	1.4859	1.6055	1.8252	1.4931	1.3914	1.4012
		0.99	9.4002	5.2929	6.3497	2.3853	4.1956	7.1966	2.7312	1.7767	2.1101
	5	0.8	1.5543	1.5014	1.5107	1.4495	1.4760	1.5277	1.4499	1.4246	1.4253
		0.9	2.8273	2.6439	2.6774	2.4669	2.5591	2.7348	2.4700	2.3854	2.3898
		0.95	5.4066	4.7444	4.8743	4.1273	4.4597	5.0699	4.1475	3.8648	3.8921
		0.99	26.1116	14.7025	17.6381	6.6258	11.6543	19.9905	7.5867	4.9351	5.8613
10	0.8	6.2173	6.0056	6.0427	5.7978	5.9041	6.1110	5.7997	5.6983	5.7011	
	0.9	11.3093	10.5755	10.7097	9.8677	10.2365	10.9391	9.8800	9.5416	9.5591	
	0.95	21.6263	18.9777	19.4972	16.5093	17.8388	20.2795	16.5901	15.4592	15.5684	
	0.99	104.4463	58.8099	70.5524	26.5030	46.6173	79.9619	30.3465	19.7401	23.4451	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 6. Estimated MSE when k=0.6, d=0.8

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1263	0.1330	0.1167	0.1191	0.1343	0.1116	0.1098	0.1100
		0.9	0.2462	0.2132	0.2354	0.1827	0.1915	0.2394	0.1682	0.1629	0.1636
		0.95	0.4682	0.3564	0.4323	0.2604	0.2941	0.4446	0.2241	0.2098	0.2131
		0.99	2.2417	0.7679	1.7822	0.0801	0.4374	1.8841	0.0698	0.0387	0.0601
	3	0.8	1.2270	1.1362	1.1971	1.0491	1.0710	1.2085	1.0029	0.9869	0.9881
		0.9	2.2158	1.9183	2.1188	1.6431	1.7228	2.1545	1.5123	1.4651	1.4712
		0.95	4.2136	3.2080	3.8903	2.3436	2.6472	4.0012	2.0164	1.8877	1.9174
		0.99	20.1751	6.9114	16.0396	0.7210	3.9366	16.9565	0.6281	0.3474	0.5407
	5	0.8	3.4082	3.1559	3.3251	2.9139	2.9747	3.3569	2.7854	2.7409	2.7440
		0.9	6.1551	5.3284	5.8855	4.5638	4.7853	5.9848	4.2004	4.0692	4.0861
		0.95	11.7045	8.9111	10.8064	6.5101	7.3532	11.1144	5.6012	5.2435	5.3261
		0.99	56.0420	19.1985	44.5544	2.0029	10.9351	47.1015	1.7448	0.9649	1.5019
10	0.8	13.6328	12.6233	13.3005	11.6547	11.8982	13.4276	11.1406	10.9622	10.9749	
	0.9	24.6204	21.3135	23.5419	18.2547	19.1408	23.9392	16.8005	16.2759	16.3434	
	0.95	46.8180	35.6446	43.2258	26.0407	29.4132	44.4577	22.4050	20.9741	16.3434	
	0.99	224.1680	76.7940	178.2178	8.0117	43.7407	188.4059	6.9793	3.8599	16.3434	
100	1	0.8	0.0622	0.0601	0.0615	0.0580	0.0585	0.0617	0.0568	0.0565	0.0565
		0.9	0.1131	0.1058	0.1107	0.0987	0.1004	0.1116	0.0949	0.0936	0.0937
		0.95	0.2163	0.1898	0.2076	0.1651	0.1721	0.2108	0.1531	0.1488	0.1493
		0.99	1.0445	0.5881	0.9008	0.2651	0.4107	0.9425	0.2025	0.1689	0.1836
	3	0.8	0.5596	0.5405	0.5532	0.5218	0.5260	0.5557	0.5112	0.5076	0.5077
		0.9	1.0178	0.9518	0.9961	0.8881	0.9037	1.0044	0.8538	0.8419	0.8427
		0.95	1.9464	1.7080	1.8685	1.4859	1.5485	1.8974	1.3778	1.3391	1.3436
		0.99	9.4002	5.2929	8.1072	2.3853	3.6959	8.4827	1.8218	1.5201	1.6520
	5	0.8	1.5543	1.5014	1.5368	1.4495	1.4611	1.5437	1.4200	1.4100	1.4103
		0.9	2.8273	2.6439	2.7668	2.4669	2.5103	2.7901	2.3715	2.3386	2.3406
		0.95	5.4066	4.7444	5.1902	4.1273	4.3015	5.2705	3.8271	3.7197	3.7322
		0.99	26.1116	14.7025	22.5199	6.6258	10.2664	23.5632	5.0604	4.2223	4.5888
10	0.8	6.2173	6.0056	6.1472	5.7978	5.8445	6.1747	5.6799	5.6399	5.6412	
	0.9	11.3093	10.5755	11.0674	9.8677	10.0413	11.1605	9.4859	9.3542	9.3624	
	0.95	21.6263	18.9777	20.7609	16.5093	17.2058	21.0822	15.3082	14.8786	14.9288	
	0.99	104.4463	58.8099	90.0797	26.5030	41.0653	94.2527	20.2414	16.8890	18.3550	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 7. Estimated MSE when k=0.7, d=0.2

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1247	0.1234	0.1137	0.1226	0.1270	0.1203	0.1117	0.1122
		0.9	0.2462	0.2083	0.2046	0.1738	0.2019	0.2157	0.1944	0.1680	0.1706
		0.95	0.4682	0.3419	0.3333	0.2361	0.3230	0.3655	0.2986	0.2211	0.2340
		0.99	2.2417	0.6731	0.7223	0.0378	0.5677	0.9122	0.3503	0.0304	0.1489
	3	0.8	1.2270	1.1221	1.1098	1.0222	1.1028	1.1427	1.0819	1.0040	1.0082
		0.9	2.2158	1.8748	1.8412	1.5634	1.8164	1.9406	1.7489	1.5109	1.5339
		0.95	4.2136	3.0769	2.9999	2.1241	2.9068	3.2895	2.6869	1.9893	2.1058
		0.99	20.1751	6.0583	6.5009	0.3396	5.1093	8.2098	3.1526	0.2729	1.3398
	5	0.8	3.4082	3.1167	3.0825	2.8391	3.0632	3.1739	3.0051	2.7883	2.8000
		0.9	6.1551	5.2076	5.1142	4.3425	5.0455	5.3905	4.8579	4.1967	4.2603
		0.95	11.7045	8.5469	8.3330	5.9004	8.0744	9.1375	7.4636	5.5258	5.8494
		0.99	56.0420	16.8287	18.0581	0.9434	14.1924	22.8051	8.7573	0.7582	3.7218
10	0.8	13.6328	12.4665	12.3294	11.3553	12.2522	12.6954	12.0197	11.1520	11.1990	
	0.9	24.6204	20.8301	20.4565	17.3694	20.1815	21.5618	19.4312	16.7859	17.0405	
	0.95	46.8180	34.1876	33.3323	23.6018	32.2976	36.5501	29.8547	22.1034	23.3978	
	0.99	224.1680	67.3149	72.2327	3.7737	56.7697	91.2205	35.0295	3.0328	14.8873	
100	1	0.8	0.0622	0.0597	0.0594	0.0573	0.0593	0.0602	0.0588	0.0569	0.0569
		0.9	0.1131	0.1046	0.1036	0.0965	0.1030	0.1063	0.1014	0.0950	0.0953
		0.95	0.2163	0.1859	0.1828	0.1579	0.1806	0.1918	0.1746	0.1531	0.1550
		0.99	1.0445	0.5425	0.5345	0.2074	0.4873	0.6296	0.3989	0.1800	0.2436
	3	0.8	0.5596	0.5374	0.5345	0.5158	0.5332	0.5418	0.5288	0.5117	0.5121
		0.9	1.0178	0.9415	0.9322	0.8682	0.9273	0.9565	0.9120	0.8547	0.8574
		0.95	1.9464	1.6727	1.6447	1.4207	1.6252	1.7257	1.5710	1.3774	1.3944
		0.99	9.4002	4.8827	4.8107	1.8662	4.3853	5.6661	3.5896	1.6196	2.1920
	5	0.8	1.5543	1.4929	1.4848	1.4327	1.4810	1.5050	1.4687	1.4212	1.4224
		0.9	2.8273	2.6151	2.5896	2.4117	2.5757	2.6569	2.5334	2.3741	2.3816
		0.95	5.4066	4.6464	4.5687	3.9463	4.5145	4.7937	4.3637	3.8262	3.8734
		0.99	26.1116	13.5630	13.3631	5.1838	12.1814	15.7392	9.9711	4.4987	6.0889
10	0.8	6.2173	5.9714	5.9393	5.7309	5.9241	6.0202	5.8750	5.6847	5.6896	
	0.9	11.3093	10.4605	10.3582	9.6467	10.3029	10.6275	10.1335	9.4961	9.5264	
	0.95	21.6263	18.5858	18.2747	15.7850	18.0579	19.1747	17.4549	15.3045	15.4934	
	0.99	104.4463	54.2520	53.4523	20.7350	48.7256	62.9567	39.8844	17.9946	24.3556	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 8. Estimated MSE when k=0.7, d=0.5

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1247	0.1281	0.1137	0.1195	0.1304	0.1140	0.1088	0.1092
		0.9	0.2462	0.2083	0.2197	0.1738	0.1927	0.2269	0.1753	0.1599	0.1617
		0.95	0.4682	0.3419	0.3811	0.2361	0.2975	0.4025	0.2438	0.2013	0.2104
		0.99	2.2417	0.6731	1.1925	0.0378	0.4511	1.3408	0.1257	0.0233	0.0790
	3	0.8	1.2270	1.1221	1.1530	1.0222	1.0749	1.1739	1.0245	0.9775	0.9807
		0.9	2.2158	1.8748	1.9774	1.5634	1.7341	2.0416	1.5763	1.4373	1.4542
		0.95	4.2136	3.0769	3.4302	2.1241	2.6776	3.6222	2.1941	1.8108	1.8928
		0.99	20.1751	6.0583	10.7325	0.3396	4.0594	12.0670	1.1309	0.2088	0.7102
	5	0.8	3.4082	3.1167	3.2026	2.8391	2.9856	3.2607	2.8455	2.7148	2.7237
		0.9	6.1551	5.2076	5.4928	4.3425	4.8167	5.6710	4.3784	3.9920	4.0391
		0.95	11.7045	8.5469	9.5284	5.9004	7.4378	10.0618	6.0946	5.0299	5.2578
		0.99	56.0420	16.8287	29.8126	0.9434	11.2761	33.5193	3.1414	<i>0.5799</i>	1.9726
10	0.8	13.6328	12.4665	12.8101	11.3553	11.9416	13.0428	11.3809	10.8577	10.8934	
	0.9	24.6204	20.8301	21.9710	17.3694	19.2663	22.6840	17.5127	15.9669	16.1552	
	0.95	46.8180	34.1876	38.1136	23.6018	29.7514	40.2472	24.3787	20.1200	21.0313	
	0.99	224.1680	67.3149	119.2504	3.7737	45.1045	134.0774	12.5656	2.3198	7.8906	
100	1	0.8	0.0622	0.0597	0.0604	0.0573	0.0586	0.0609	0.0574	0.0562	0.0563
		0.9	0.1131	0.1046	0.1071	0.0965	0.1008	0.1088	0.0967	0.0928	0.0931
		0.95	0.2163	0.1859	0.1950	0.1579	0.1731	0.2008	0.1589	0.1463	0.1477
		0.99	1.0445	0.5425	0.7055	0.2074	0.4192	0.7726	0.2509	0.1484	0.1879
	3	0.8	0.5596	0.5374	0.5438	0.5158	0.5269	0.5484	0.5160	0.5055	0.5059
		0.9	1.0178	0.9415	0.9639	0.8682	0.9066	0.9792	0.8697	0.8350	0.8371
		0.95	1.9464	1.6727	1.7548	1.4207	1.5578	1.8068	1.4302	1.3164	1.3290
		0.99	9.4002	4.8827	6.3497	1.8662	3.7726	6.9538	2.2581	1.3349	1.6907
	5	0.8	1.5543	1.4929	1.5107	1.4327	1.4636	1.5234	1.4334	1.4041	1.4051
		0.9	2.8273	2.6151	2.6774	2.4117	2.5184	2.7201	2.4158	2.3193	2.3251
		0.95	5.4066	4.6464	4.8743	3.9463	4.3272	5.0190	3.9727	3.6565	3.6916
		0.99	26.1116	13.5630	17.6381	5.1838	10.4795	19.3161	6.2725	3.7079	4.6962
10	0.8	6.2173	5.9714	6.0427	5.7309	5.8543	6.0937	5.7335	5.6165	5.6203	
	0.9	11.3093	10.4605	10.7097	9.6467	10.0734	10.8805	9.6631	9.2770	9.3001	
	0.95	21.6263	18.5858	19.4972	15.7850	17.3088	20.0761	15.8906	14.6258	14.7662	
	0.99	104.4463	54.2520	70.5524	20.7350	41.9180	77.2645	25.0898	14.8312	18.7847	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 9. Estimated MSE when k=0.7, d=0.8

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1247	0.1330	0.1137	0.1166	0.1340	0.1080	0.1061	0.1062
		0.9	0.2462	0.2083	0.2354	0.1738	0.1843	0.2384	0.1581	0.1523	0.1532
		0.95	0.4682	0.3419	0.4323	0.2361	0.2751	0.4413	0.1991	0.1841	0.1881
		0.99	2.2417	0.6731	1.7822	0.0378	0.3674	1.8533	0.0389	0.0188	0.0335
	3	0.8	1.2270	1.1221	1.1971	1.0222	1.0481	1.2056	0.9703	0.9522	0.9537
		0.9	2.2158	1.8748	2.1188	1.5634	1.6574	2.1453	1.4212	1.3691	1.3768
		0.95	4.2136	3.0769	3.8903	2.1241	2.4752	3.9715	1.7915	1.6562	1.6918
		0.99	20.1751	6.0583	16.0396	0.3396	3.3064	16.6799	0.3493	0.1685	0.3004
	5	0.8	3.4082	3.1167	3.3251	2.8391	2.9110	3.3488	2.6947	2.6443	2.6484
		0.9	6.1551	5.2076	5.8855	4.3425	4.6037	5.9590	3.9473	3.8026	3.8240
		0.95	11.7045	8.5469	10.8064	5.9004	6.8757	11.0321	4.9764	4.6005	4.6993
		0.99	56.0420	16.8287	44.5544	0.9434	9.1844	46.3331	0.9702	0.4681	0.8345
10	0.8	13.6328	12.4665	13.3005	11.3553	11.6432	13.3951	10.7776	10.5756	10.5922	
	0.9	24.6204	20.8301	23.5419	17.3694	18.4143	23.8360	15.7880	15.2092	15.2949	
	0.95	46.8180	34.1876	43.2258	23.6018	27.5030	44.1283	19.9058	18.4022	18.7972	
	0.99	224.1680	67.3149	178.2178	3.7737	36.7379	185.3324	3.8809	1.8724	3.3380	
100	1	0.8	0.0622	0.0597	0.0615	0.0573	0.0579	0.0617	0.0560	0.0556	0.0556
		0.9	0.1131	0.1046	0.1107	0.0965	0.0985	0.1114	0.0922	0.0907	0.0908
		0.95	0.2163	0.1859	0.2076	0.1579	0.1661	0.2100	0.1448	0.1400	0.1406
		0.99	1.0445	0.5425	0.9008	0.2074	0.3647	0.9307	0.1559	0.1247	0.1398
	3	0.8	0.5596	0.5374	0.5532	0.5158	0.5207	0.5551	0.5037	0.4995	0.4997
		0.9	1.0178	0.9415	0.9961	0.8682	0.8867	1.0023	0.8295	0.8160	0.8170
		0.95	1.9464	1.6727	1.8685	1.4207	1.4947	1.8899	1.3023	1.2595	1.2652
		0.99	9.4002	4.8827	8.1072	1.8662	3.2818	8.3766	1.4023	1.1214	1.2578
	5	0.8	1.5543	1.4929	1.5368	1.4327	1.4465	1.5419	1.3989	1.3874	1.3879
		0.9	2.8273	2.6151	2.7668	2.4117	2.4630	2.7842	2.3040	2.2665	2.2692
		0.95	5.4066	4.6464	5.1902	3.9463	4.1518	5.2497	3.6175	3.4984	3.5144
		0.99	26.1116	13.5630	22.5199	5.1838	9.1161	23.2683	3.8951	3.1149	3.4939
10	0.8	6.2173	5.9714	6.1472	5.7309	5.7859	6.1677	5.5957	5.5496	5.5514	
	0.9	11.3093	10.4605	11.0674	9.6467	9.8519	11.1367	9.2156	9.0659	9.0767	
	0.95	21.6263	18.5858	20.7609	15.7850	16.6072	20.9990	14.4698	13.9932	14.0574	
	0.99	104.4463	54.2520	90.0797	20.7350	36.4642	93.0733	15.5802	12.4593	13.9752	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 10. Estimated MSE when k=0.9, d=0.2

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1217	0.1234	0.1081	0.1191	0.1246	0.1162	0.1056	0.1064
		0.9	0.2462	0.1992	0.2046	0.1575	0.1915	0.2082	0.1823	0.1508	0.1546
		0.95	0.4682	0.3154	0.3333	0.1937	0.2941	0.3434	0.2651	0.1783	0.1958
		0.99	2.2417	0.5301	0.7223	0.0227	0.4374	0.7761	0.2297	0.0168	0.0854
	3	0.8	1.2270	1.0947	1.1098	0.9706	1.0710	1.1205	1.0447	0.9484	0.9550
		0.9	2.2158	1.7923	1.8412	1.4156	1.7228	1.8733	1.6395	1.3553	1.3892
		0.95	4.2136	2.8383	2.9999	1.7424	2.6472	3.0910	2.3858	1.6035	1.7613
		0.99	20.1751	4.7705	6.5009	0.2033	3.9366	6.9844	2.0674	0.1504	0.7679
	5	0.8	3.4082	3.0407	3.0825	2.6954	2.9747	3.1124	2.9016	2.6339	2.6522
		0.9	6.1551	4.9784	5.1142	3.9317	4.7853	5.2034	4.5538	3.7640	3.8584
		0.95	11.7045	7.8841	8.3330	4.8400	7.3532	8.5861	6.6273	4.4542	4.8924
		0.99	56.0420	13.2514	18.0581	0.5647	10.9351	19.4012	5.7429	0.4178	2.1329
10	0.8	13.6328	12.1622	12.3294	10.7802	11.8982	12.4492	11.6056	10.5338	10.6070	
	0.9	24.6204	19.9132	20.4565	15.7255	19.1408	20.8131	18.2145	15.0550	15.4324	
	0.95	46.8180	31.5367	33.3323	19.3600	29.4132	34.3446	26.5095	<i>17.8169</i>	19.5697	
	0.99	224.1680	53.0058	72.2327	2.2591	43.7407	77.6050	22.9718	<i>1.6713</i>	8.5317	
100	1	0.8	0.0622	0.0591	0.0594	0.0561	0.0585	0.0597	0.0578	0.0555	0.0556
		0.9	0.1131	0.1024	0.1036	0.0923	0.1004	0.1045	0.0983	0.0904	0.0909
		0.95	0.2163	0.1784	0.1828	0.1444	0.1721	0.1857	0.1646	0.1388	0.1416
		0.99	1.0445	0.4662	0.5345	0.1233	0.4107	0.5630	0.3144	0.1036	0.1730
	3	0.8	0.5596	0.5314	0.5345	0.5040	0.5260	0.5369	0.5204	0.4988	0.4995
		0.9	1.0178	0.9213	0.9322	0.8299	0.9037	0.9402	0.8845	0.8133	0.8176
		0.95	1.9464	1.6055	1.6447	1.2988	1.5485	1.6709	1.4813	1.2486	1.2740
		0.99	9.4002	4.1956	4.8107	1.1091	3.6959	5.0670	2.8297	0.9318	1.5563
	5	0.8	1.5543	1.4760	1.4848	1.3999	1.4611	1.4915	1.4455	1.3855	1.3874
		0.9	2.8273	2.5591	2.5896	2.3051	2.5103	2.6116	2.4569	2.2591	2.2710
		0.95	5.4066	4.4597	4.5687	3.6077	4.3015	4.6415	4.1146	3.4681	3.5388
		0.99	26.1116	11.6543	13.3631	3.0808	10.2664	14.0751	7.8603	2.5881	4.3229
10	0.8	6.2173	5.9041	5.9393	5.5996	5.8445	5.9660	5.7819	5.5417	5.5496	
	0.9	11.3093	10.2365	10.3582	9.2202	10.0413	10.4465	9.8275	9.0362	9.0838	
	0.95	21.6263	17.8388	18.2747	14.4306	17.2058	18.5659	16.4584	13.8722	14.1550	
	0.99	104.4463	46.6173	53.4523	12.3231	41.0653	56.3003	31.4412	10.3522	17.2912	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 11. Estimated MSE when n=100, k=0.9, d=0.5

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1217	0.1281	0.1081	0.1153	0.1289	0.1085	0.1022	0.1027
		0.9	0.2462	0.1992	0.2197	0.1575	0.1808	0.2220	0.1596	0.1416	0.1443
		0.95	0.4682	0.3154	0.3811	0.1937	0.2662	0.3879	0.2046	0.1585	0.1705
		0.99	2.2417	0.5301	1.1925	0.0227	0.3386	1.2353	0.0595	0.0124	0.0362
	3	0.8	1.2270	1.0947	1.1530	0.9706	1.0369	1.1598	0.9742	0.9168	0.9217
		0.9	2.2158	1.7923	1.9774	1.4156	1.6262	1.9982	1.4351	1.2721	1.2969
		0.95	4.2136	2.8383	3.4302	1.7424	2.3957	3.4910	1.8403	1.4257	1.5336
		0.99	20.1751	4.7705	10.7325	0.2033	3.0474	11.1181	0.5349	0.1103	0.3246
	5	0.8	3.4082	3.0407	3.2026	2.6954	2.8800	3.2217	2.7055	2.5457	2.5594
		0.9	6.1551	4.9784	5.4928	3.9317	4.5169	5.5505	3.9859	3.5330	3.6018
		0.95	11.7045	7.8841	9.5284	4.8400	6.6547	9.6973	5.1120	3.9601	4.2600
		0.99	56.0420	13.2514	29.8126	0.5647	8.4651	30.8837	1.4857	0.3064	0.9014
10	0.8	13.6328	12.1622	12.8101	10.7802	11.5189	12.8864	10.8204	10.1808	10.2359	
	0.9	24.6204	19.9132	21.9710	15.7255	18.0671	22.2019	15.9424	14.1306	14.4057	
	0.95	46.8180	31.5367	38.1136	19.3600	26.6191	38.7893	20.4483	15.8405	17.0400	
	0.99	224.1680	53.0058	119.2504	2.2591	33.8606	123.5349	5.9429	1.2255	3.6058	
100	1	0.8	0.0622	0.0591	0.0604	0.0561	0.0576	0.0606	0.0561	0.0547	0.0547
		0.9	0.1131	0.1024	0.1071	0.0923	0.0976	0.1077	0.0925	0.0878	0.0882
		0.95	0.2163	0.1784	0.1950	0.1444	0.1632	0.1969	0.1460	0.1310	0.1331
		0.99	1.0445	0.4662	0.7055	0.1233	0.3446	0.7259	0.1731	0.0822	0.1227
	3	0.8	0.5596	0.5314	0.5438	0.5040	0.5181	0.5454	0.5044	0.4912	0.4917
		0.9	1.0178	0.9213	0.9639	0.8299	0.8784	0.9689	0.8322	0.7895	0.7927
		0.95	1.9464	1.6055	1.7548	1.2988	1.4688	1.7717	1.3133	1.1787	1.1974
		0.99	9.4002	4.1956	6.3497	1.1091	3.1008	6.5331	1.5575	0.7389	1.1035
	5	0.8	1.5543	1.4760	1.5107	1.3999	1.4392	1.5149	1.4010	1.3643	1.3658
		0.9	2.8273	2.5591	2.6774	2.3051	2.4399	2.6915	2.3116	2.1928	2.2018
		0.95	5.4066	4.4597	4.8743	3.6077	4.0800	4.9213	3.6480	3.2741	3.3260
		0.99	26.1116	11.6543	17.6381	3.0808	8.6133	18.1475	4.3263	2.0524	3.0650
10	0.8	6.2173	5.9041	6.0427	5.5996	5.7569	6.0596	5.6038	5.4568	5.4628	
	0.9	11.3093	10.2365	10.7097	9.2202	9.7593	10.7658	9.2462	8.7711	8.8071	
	0.95	21.6263	17.8388	19.4972	14.4306	16.3200	19.6851	14.5917	13.0962	13.3036	
	0.99	104.4463	46.6173	70.5524	12.3231	34.4532	72.5899	17.3050	8.2091	12.2596	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 12. Estimated MSE when n=100, k=0.9, d=0.8

N	σ^2	rho	OLS	RIDGE	LIU	K-L	MRT	TP	NTPKL	MTPKL1	MTPKL2
50	1	0.8	0.1363	0.1217	0.1330	0.1081	0.1118	0.1333	0.1013	0.0989	0.0992
		0.9	0.2462	0.1992	0.2354	0.1575	0.1710	0.2364	0.1399	0.1333	0.1345
		0.95	0.4682	0.3154	0.4323	0.1937	0.2422	0.4351	0.1578	0.1421	0.1471
		0.99	2.2417	0.5301	1.7822	0.0227	0.2703	1.8030	0.0131	0.0101	0.0117
	3	0.8	1.2270	1.0947	1.1971	0.9706	1.0046	1.1999	0.9087	0.8867	0.8890
		0.9	2.2158	1.7923	2.1188	1.4156	1.5378	2.1274	1.2567	1.1968	1.2079
		0.95	4.2136	2.8383	3.8903	1.7424	2.1794	3.9162	1.4186	1.2771	1.3226
		0.99	20.1751	4.7705	16.0396	0.2033	2.4324	16.2272	0.1167	0.0898	0.1042
	5	0.8	3.4082	3.0407	3.3251	2.6954	2.7899	3.3329	2.5232	2.4621	2.4685
		0.9	6.1551	4.9784	5.8855	3.9317	4.2714	5.9094	3.4902	3.3237	3.3545
		0.95	11.7045	7.8841	10.8064	4.8400	6.0539	10.8783	3.9406	3.5473	3.6738
		0.99	56.0420	13.2514	44.5544	0.5647	6.7567	45.0756	0.3240	0.2492	0.2892
10	0.8	13.6328	12.1622	13.3005	10.7802	11.1585	13.3316	10.0908	9.8463	9.8718	
	0.9	24.6204	19.9132	23.5419	15.7255	17.0847	23.6374	13.9592	13.2930	13.4163	
	0.95	46.8180	31.5367	43.2258	19.3600	24.2157	43.5134	15.7624	14.1894	14.6951	
	0.99	224.1680	53.0058	178.2178	2.2591	27.0269	180.3023	1.2961	0.9966	1.1570	
100	1	0.8	0.0622	0.0591	0.0615	0.0561	0.0568	0.0615	0.0544	0.0539	0.0539
		0.9	0.1131	0.1024	0.1107	0.0923	0.0949	0.1109	0.0871	0.0853	0.0855
		0.95	0.2163	0.1784	0.2076	0.1444	0.1551	0.2084	0.1295	0.1240	0.1249
		0.99	1.0445	0.4662	0.9008	0.1233	0.2934	0.9100	0.0926	0.0670	0.0816
	3	0.8	0.5596	0.5314	0.5532	0.5040	0.5105	0.5539	0.4889	0.4838	0.4840
		0.9	1.0178	0.9213	0.9961	0.8299	0.8541	0.9981	0.7832	0.7667	0.7682
		0.95	1.9464	1.6055	1.8685	1.2988	1.3954	1.8755	1.1648	1.1149	1.1233
		0.99	9.4002	4.1956	8.1072	1.1091	2.6407	8.1898	0.8325	0.6026	0.7334
	5	0.8	1.5543	1.4760	1.5368	1.3999	1.4179	1.5385	1.3579	1.3436	1.3443
		0.9	2.8273260	2.5591	2.7668	2.3051	2.3725	2.7726	2.1753	2.1296	2.1338
		0.95	5.4066	4.4597	5.1902	3.6077	3.8759	5.2096	3.2353	3.0967	3.1201
		0.99	26.1116	11.6543	22.5199	3.0808	7.3351	22.7494	2.3123	1.6738	2.0371
10	0.8	6.2173	5.9041	6.1472	5.5996	5.6715	6.1540	5.4315	5.3740	5.3768	
	0.9	11.3093	10.2365	11.0674	9.2202	9.4898	11.0902	8.7008	8.5182	8.5349	
	0.95	21.6263	17.8388	20.7609	14.4306	15.5034	20.8384	12.9409	12.3866	12.4801	
	0.99	104.4463	46.6173	90.0797	12.3231	29.3403	90.9977	9.2487	6.6946	8.1480	

Two estimators with Minimum MSE values are bolded in each row. The smaller of the two is italicized.

Table 13. The results of regression coefficients and the corresponding MSE values.

	$\hat{\alpha}$	$\hat{\alpha}_{RE}$	$\hat{\alpha}_{LE}$	$\hat{\alpha}_{KL}$	$\hat{\alpha}_{NTP}$	$\hat{\alpha}_{MRT}$	$\hat{\alpha}_{NTPKL}$	$\hat{\alpha}_{MTPKL1}$	$\hat{\alpha}_{MTPKL2}$
$\hat{\alpha}_0$	-52.9936	-49.3576	-52.9936	-45.7216	41.6554	41.6554	-24.8522	-49.3576	-52.9936
$\hat{\alpha}_1$	0.0711	0.0703	0.0711	0.0696	47.4077	47.40767	47.42099	0.070317	0.071073
$\hat{\alpha}_2$	-0.4142	-0.4079	-0.4142	-0.4015	13.7557	13.75574	13.64773	-0.40785	-0.4142
$\hat{\alpha}_3$	-0.4235	-0.4323	-0.4235	-0.4411	0.8644	0.864403	1.018985	-0.43227	-0.42347
$\hat{\alpha}_4$	-0.5726	-0.5746	-0.5726	-0.5767	0.3746	0.374561	0.415282	-0.57463	-0.57257
$\hat{\alpha}_5$	48.4179	48.0144	48.4179	47.6109	-4.7292	-4.72919	2.27232	48.01439	48.41786
MSE	17095.15	15027.25	17095.14	13122.26	17095.14	13316.98	11510.3	2915.66	5150.228